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*Published in:*  
 Movement Disorders Clinical Practice

*DOI:*  
[10.1002/mdc3.12339](https://doi.org/10.1002/mdc3.12339)

**IMPORTANT NOTE: You are advised to consult the publisher's version (publisher's PDF) if you wish to cite from it. Please check the document version below.**

*Document Version*  
 Publisher's PDF, also known as Version of record

*Publication date:*  
 2016

[Link to publication in University of Groningen/UMCG research database](#)

*Citation for published version (APA):*

Kuiper, M. J., Vrijenhoek, L., Brandsma, R., Lunsing, R. J., Burger, H., Eggink, H., Peall, K. J., Contarino, M. F., Speelman, J. D., Tijssen, M. A. J., & Sival, D. A. (2016). The Burke-Fahn-Marsden Dystonia Rating Scale is Age-Dependent in Healthy Children. *Movement Disorders Clinical Practice*, 3(6), 580-586. <https://doi.org/10.1002/mdc3.12339>

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# The Burke-Fahn-Marsden Dystonia Rating Scale is Age-Dependent in Healthy Children

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**Abstract:** Background: The Burke-Fahn-Marsden Dystonia Rating Scale is a universally applied instrument for the quantitative assessment of dystonia in both children and adults. However, immature movements by healthy young children may also show “dystonic characteristics” as a consequence of physiologically incomplete brain maturation. This could implicate that Burke-Fahn-Marsden scale scores are confounded by pediatric age.

Objective: In healthy young children, we aimed to determine whether physiologically immature movements and postures can induce an age-related effect on Burke-Fahn-Marsden movement and disability scale scores.

Methods: Nine assessors specialized in movement disorders (3 adult neurologists, 3 pediatric neurologists, and 3 MD/PhD students) independently scored the Burke-Fahn-Marsden movement scale in 52 healthy children (4–16 years of age; 2 boys and 2 girls per year of age). Independent of that, parents scored their children’s functional motor development according to the Burke-Fahn-Marsden disability scale in another 52 healthy children (4–16 years of age; 2 boys and 2 girls per year of age). By regression analysis, we determined the association between Burke-Fahn-Marsden movement and disability scales outcomes and pediatric age.

Results: In healthy children, assessment of physiologically immature motor performances by the Burke-Fahn-Marsden movement and disability scales showed an association between the outcomes of both scales and age (until 16 years and 12 years of age,  $\beta = -0.72$  and  $\beta = -0.60$ , for Burke-Fahn-Marsden movement and disability scale, respectively [both  $P < 0.001$ ]).

Conclusions: The Burke-Fahn-Marsden movement and disability scales are influenced by the age of the child. For accurate interpretation of longitudinal Burke-Fahn-Marsden Dystonia Rating Scale scores in young dystonic children, consideration of pediatric age-relatedness appears advisory.

Dystonia is a movement disorder characterized by sustained or intermittent muscle contractions, causing abnormal, often repetitive, movements or postures.<sup>1–3</sup> The neuroanatomical substrate for dystonia is ascribed to dysfunctional networks of the basal ganglia, cerebellum, thalamus, cerebral cortex, and brainstem.<sup>4</sup> The

term *early onset dystonia* is used to denounce the initiation of dystonia before the twenty-sixth year of life.<sup>1</sup> As the characterization spans distinctly different developmental stages, pediatric subdivision into subgroups of infancy (0–2 years), childhood (3–12 years), and adolescence (13–20 years) has been advocated.<sup>1</sup>

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**Keywords:** Burke-Fahn-Marsden Dystonia Rating Scale, child, dystonia, validation, age.

#Marieke Kuiper and Lois Vrijenhoek contributed equally to this study.

Relevant disclosures and conflicts of interest are listed at the end of this article.

Received 1 August 2015; revised 23 December 2015; accepted 11 January 2016.

Published online 3 May 2016 in Wiley InterScience (www.interscience.wiley.com). DOI:10.1002/mdc3.12339

The Burke-Fahn-Marsden Dystonia Rating Scale (BFMDRS) is a universally applied biomarker for the severity of dystonia. The scale consists of a movement and disability subscale (Burke-Fahn-Marsden Movement Scale [BFMMS] and Burke-Fahn-Marsden Disability Scale [BFMDS], respectively).<sup>5</sup> The BFMMS measures dystonia in nine body regions (including the eyes, mouth, speech and swallowing, neck, trunk, arms, and legs) with scores ranging from 0 (minimum) to 120 (maximum). The BFMDRS is a functional marker consisting of parental- or self-reported daily activities (involving speech, handwriting, feeding, eating, swallowing, hygiene, dressing, and walking), with scores ranging from 0 (completely independent) to 30 (completely dependent). Although BFMDRS was originally developed as an instrument for the measurement of primary torsion dystonia in adults, the scale is now uniformly being applied to quantify dystonia severity in children as well.<sup>6</sup>

In healthy young children, it was demonstrated that incomplete maturation of pediatric cerebral networks (involving the basal ganglia, cerebellum, brainstem, and cortex)<sup>4,7–13</sup> is reflected by developmental movements and postures.<sup>6,14–22</sup> These physiological, immature movements and postures can transiently show features that fulfill the criteria for “dystonia” or “ataxia” (such as the asymmetrical tonic neck reflex before 6 months of age<sup>17</sup> or the scissoring grasp in toddlers<sup>14</sup>).<sup>6,14–17</sup> Complex motor tasks by healthy school children may also show dystonic characteristics such as during writing, playing the piano, finger or foot tapping, or the fog test.<sup>18,19</sup> Because these physiological features are attributed to incomplete maturation of the central nervous system, they are likely to disappear when the child matures.<sup>6,20–23</sup> For adequate interpretation of longitudinal BFMDRS scores from pediatric to adult age, this would thus implicate that one may first need to consider the effect by age (i.e., by physiologic cerebral maturation) on the scores.

In a large cohort of healthy children, we therefore aimed to evaluate the influence of age on BFMDRS (BFMMS and BFMDRS) scores. To our best knowledge, BFMDRS scores have never been studied for potential age-relatedness in children before. We reasoned that forthcoming insight in potential BFMDRS age dependency could provide information for: (1) reliable longitudinal treatment evaluation in young children (such as for longitudinal dystonia databases and for longitudinal evaluation of innovative therapies (such as deep brain stimulation [DBS])<sup>24,25</sup>; (2) understanding of dystonia progression in different “age-of-onset” groups<sup>1</sup>; and (3) adequate phenotypic discrimination between “immature” and “dystonic” motor patterns for adequate interpretation of next-generation sequencing (NGS) panels.<sup>3,26</sup>

## Methods

### Participants

After informed consent by the parents and children (when older than 12 years of age), we included a total of 104 healthy children for the investigation of BFMDRS age-relatedness. In the

absence of existing quantitative age-related BFMDRS outcomes in children, we based sample size selection on previously published data on interobserver agreement in dystonic children.<sup>27</sup> Detecting an Intraclass Correlation Coefficient (ICC) of 0.80 for the total score or over the null hypothesis of a moderate ICC of 0.60 (0.86 published for children),<sup>27</sup> a sample size of 36 children would be needed. Using a significance level ( $\alpha$ ) of 0.05 would imply that inclusion of 52 children would be sufficient.

For the investigation of potential BFMDRS age-relatedness, we thus included 104 healthy children (4–16 years of age; 2 boys and 2 girls per year of age,  $n = 52$  children for each BFMMS and BFMDRS subscale), following mainstream education at school. Before decision on study inclusion, the parents of the child completed a detailed questionnaire concerning the health of their child. This questionnaire involved neurological and/or skeletal diagnoses, prescribed medication, school performances, sporting activities, and parental education level. Participants were excluded from the study if they: (1) were diagnosed with a neurological or skeletal disorder; (2) showed a positive Gower’s sign; (3) received medication with known side-effects on motor behavior; (4) presented with developmental delay or cognitive impairment imposing the need for extra support by special schools. We recruited participants by open advertisements at regional schools. Analogous to previous age validation studies of ataxia rating scales,<sup>21</sup> we did not exclude pediatric behavioral diagnoses such as Attention Deficit Hyperactive Disorder (ADHD) or Attention Deficit Disorder (ADD). For subject characteristics, see Table S1.

## Procedure

The study was approved by the medical ethical committee of the University Medical Center Groningen, the Netherlands. Collected physiognomic data included length, weight, and head circumference.

## Test and Scoring Methods

### 1. BFMMS

In a set of 52 healthy children (4–16 years of age; 2 boys and 2 girls per year of age), we video-recorded BFMMS in a quiet place, in accordance with a standardized video protocol (see Table S2).<sup>28</sup> Nine independent assessors from the movement-disorder team (involving experienced pediatric neurologists [ $n = 3$ ] and adult neurologists [ $n = 3$ ] and less experienced [but weekly trained] MD/PhD research students [ $n = 3$ ]) scored the BFMMS video recordings offline. Before the study, pilot data on BFMMS interobserver agreement in dystonic children (scored by one pediatric neurologist [D.A.S.], one adult neurologist [M.A.J.T.], and one MD/PhD student [H.E.]) showed appropriate interobserver agreement (ICC > 0.90, i.e., excellent when interpreted according to Cicchetti<sup>29</sup> and almost perfect when interpreted according to Landis and Koch<sup>30</sup>). All assessors

received the same protocol and the written information indicating that they should assess the children's motor behavior according to the definition of dystonia,<sup>1</sup> following BFMMS instructions,<sup>5</sup> identical to the way they would assess the same motor behavior in an adult patient. The assessors were aware that their BFMMS scores should not include other immature, developmental features that do not fulfill criteria for dystonia. We determined the mean outcome of nine assessments per child, resulting in four data points per year of age (in the age range of 4–16 years). We subsequently associated the mean BFMMS scores with age and we determined inter- and intraobserver agreement and test–retest reliability (after a latent time interval of more than 3 weeks).

## 2. BFMDS

In a second set of 52 healthy children (4–16 years of age; 2 boys and 2 girls per year of age), we obtained BFMDS scores by parental reports on their children's performances. We thus obtained four data points per year of age (in the age range of 4–16 years). We subsequently associated BFMDS scores with age.

## 3. BFMMS and BFMDS

We compared the age-dependency of mean total BFMMS scores and BFMDS scores.

## Statistical Analysis

We performed statistical analyses using PASW Statistics 20 for Windows (SPSS Inc., Chicago IL). We assessed normality of the distribution of the BFMMS and BFMDS total score, both graphically and using the Kolmogorov-Smirnov test. With multivariable regression analysis we analyzed the influence of age, gender, school performances, sporting activities, and parental education level on BFMMS and BFMDS scores. When the variables significantly influenced the model, we calculated unstandardized (B) and standardized ( $\beta$ ) regression coefficients. We examined outliers'  $\geq 3$  SD in more detail by calculating DF $\beta$ . When outliers were present (DF $\beta > 1$ ), they were removed from the regression model. We also performed logarithmic analysis to assemble the best-fitted trend line.

To check for the reliability of BFMMS scores, we determined inter- and intraobserver agreement and test–retest reliability of BFMMS outcomes by ICCs, using the two-way mixed model and single measurement coefficients. According to Cicchetti,<sup>29</sup> official cutoffs for qualitative rating of ICC values are as follows: ICC < 0.40: poor; 0.40 to 0.59: fair; 0.60 to 0.74: good; 0.75 to 1.00: excellent. For uniformity reasons with previously published data,<sup>21,22</sup> we also interpreted outcomes by Landis and Koch criteria,<sup>30</sup> which are originally described for categorical data. According to Landis and Koch<sup>30</sup> we characterized ICC outcomes as follows: ICC < 0.20: slight; 0.21 to 0.40: fair; 0.41 to 0.60: moderate; 0.61 to 0.80: substantial; >0.81: almost perfect.

To compare the age-dependency of the BFMMS and BFMDS scores, we fitted two linear regression models in the first and second set of children, respectively. To allow meaningful comparison between the age-dependency of the scales, the scores were transformed into z scores in advance of the analyses. Subsequently we calculated the difference between the two regression coefficients and tested its statistical significance using the Z-test.<sup>31</sup>

All statistical tests were two-sided. The *P* values of <0.05 (two-sided) were considered as statistically significant.

## Results

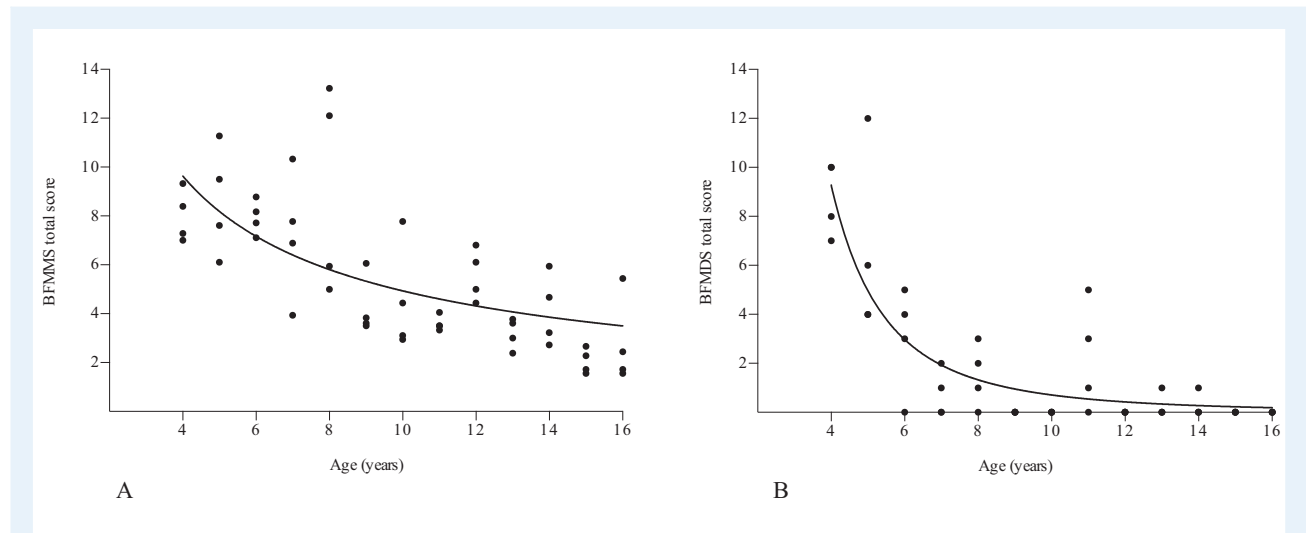
### Characteristics of Included Children

At inclusion, only a minority of the children was diagnosed with a medical condition (involving asthma [*n* = 2], bowel problems [*n* = 1], and ADHD [*n* = 1]). For patient characteristics, see Table S1. The included children participated more frequently in sports (48–69% > 2 h/week) than the average Dutch population.<sup>32</sup> Most parents of included children (58–69%) had achieved academic grades, compared to a minority (26–30%) of the average Dutch population.<sup>33</sup> Additionally, only 4% of the included children exhibited school performances below average, compared to 25% of the average Dutch population.<sup>33</sup>

#### 1. BFMMS

**Total BFMMS Scores.** Total BFMMS scores were not normally distributed (*P* < 0.05). The criteria of multivariable linear regression were met. Total BFMMS scores were significantly predicted by age, both for the total observer group ( $\beta = -0.72$ ; *P* < 0.001), as for the three observer subgroups (pediatric neurologists:  $\beta = -0.64$ ; *P* < 0.001; adult neurologists:  $\beta = -0.64$ ; *P* < 0.001; research students:  $\beta = -0.57$ ; *P* < 0.001). Age explained 51.9% of difference in scores of the total observer group (*P* < 0.001). Effects of gender, sporting activities, school performances, and the educational level of the parents did not significantly influence BFMMS scores (see Table S3). As age was the only significant predictor for total BFMMS scores, the inverse relation between age and mean total scores was determined by logarithmic analysis with a logarithmic trend (log–log line). The consistency between age and mean total BFMMS scores showed an age-related effect until 16 years of age (Fig. 1A).

**BFMMS Subscale Scores.** Subscale scores were not normally distributed (*P* < 0.05). The “arms” subscales (i.e., during pronation and supination, finger tapping, writing/drawing, and/or at rest) showed the highest scores in comparison with the other subscales (*P* < 0.001). Legs, trunk, and mouth also contributed significantly to the total scores (*P* < 0.01). With multivariable linear regression analysis we observed a significant, inverse age-effect on the subscale scores of the arms, legs, mouth, and



**Figure 1** BFMS (A) and BFMD (B) scores related to age. Data points represent (mean) individual scores per child. BFMS and BFMD are age-dependent until 16 and 12 years of age, respectively. BFMS, Burke-Fahn-Marsden Movement Scale; BFMD, Burke-Fahn-Marsden Disability Scale.

trunk ( $P < 0.05$ ). An age-relationship was absent for the subscale scores of the eyes, speech and swallowing, and neck. For video examples of age-related BFMS performances, see the included video recordings.

*Observer Agreement and Test–Retest Reliability of BFMS.* Interobserver agreement showed statistically significant ICCs of 0.40 (total group,  $P < 0.001$ ), 0.62 (pediatric neurologists,  $P < 0.001$ ), 0.24 (adult neurologists,  $P = 0.002$ ), and 0.47 (research students,  $P < 0.001$ ). ICC outcomes revealed a “fair” interobserver agreement for the total group (according to criteria from both Cicchetti<sup>29</sup> and Landis and Koch<sup>30</sup>). The ICC for the three individual observer-subgroups varied between “poor to good” and/or “fair to substantial” (according to both Cicchetti’s<sup>29</sup> and Landis’s and Koch’s<sup>30</sup> criteria, respectively). For further information on inter- and intraobserver agreement and test–retest reliability, see Table S4.

## 2. BFMD

*Total BFMD Scores.* Total BFMD scores were not normally distributed ( $P < 0.001$ ). The criteria of multivariable linear regression were met. Total BFMD scores were significantly predicted by age ( $\beta = -0.60$ ;  $P < 0.001$ ). Age explained 36.2% of difference in scores ( $P < 0.001$ ). Effects of gender, sporting activities, school performances, and the educational level of the parents did not significantly influence BFMD scores (see Table S3). As age was the only significant predictor for total BFMD scores, the inverse relation between age and total BFMD scores was determined by logarithmic analysis with a logarithmic trend (log–log line). The consistency between age and total BFMD scores showed an age-related effect until about 12 years of age (Fig. 1B).

## Association between BFMS and BFMD Age-Dependency

The unstandardized regression coefficients between BFMS ( $B = -0.19$ ) and BFMD age-dependency ( $B = -0.18$ ) revealed no significant difference ( $P = 0.75$ ).

## Discussion

In healthy children, we investigated the potential influence by age on BFMDRS scores. Our results indicate that both movement and disability subscales (i.e., BFMS and BFMD) are influenced by age (until 16 and 12 years of age, respectively). Additionally, BFMS and BFMD scores showed a similar pattern of age-dependency, suggesting that BFMS age-relatedness has functional implications.

The present study indicates that BFMS scores are age-dependent, at least until 16 years of age. Significantly, this result was provided by all observer-subgroups, that is, by the total-observer group, by the experienced pediatric and adult neurologists, and also by the less-experienced research students. Because dystonic pediatric outcome parameters should be interpretable against healthy reference values, this would implicate that insight in pediatric age-related BFMS reference values would be needed first. When measured against the theoretically maximal BFMS score, it appears that the quantitative age-related BFMS effect seems relatively small. However, because young children are often at an early disease stage (i.e., remote from the theoretical maximum), consideration of the age-related effect appears advisory. Furthermore, BFMS and BFMD scores showed a similar age-related effect, suggesting that the BFMS age-relatedness is also reflected by functional changes. Because functional assessments are increasingly advocated as the best

treatment outcome parameters,<sup>34,35</sup> we would therefore suggest that both BFMDRS subscales should be interpreted for age. Such data may appear of special interest for longitudinal treatment trials in small, heterogeneous groups of dystonic children, measuring relatively small effects throughout time (such as for dystonic children receiving deep brain stimulation at an increasingly younger age). Under intraindividually, longitudinally assessed conditions, small quantitative changes in BFMMS and BFMDS scores could thus run the theoretical risk of being overinterpreted, as “therapeutically” and/or “functionally” effective.<sup>36</sup>

Interestingly, the BFMMS age-dependency lingered until 16 years of age, revealing no optimum score (zero) in the oldest included children (16 years of age). However, in the absence of reference values in healthy adults, it is tempting to speculate that BFMMS scores of 16-year-old children are likely to approach adult optimality as a reflection of physiologic brain maturation. The basal ganglia receive signals from several cortical areas (i.e., motor, somatosensory and [pre]frontal cortex, and the limbic system) and the cerebellum. They modulate and transport these signals via the thalamus, to the brainstem, cortical motor areas (such as primary motor, premotor, and oculomotor cortex), posterior parietal cortex, and the temporal cortex.<sup>7–9</sup> Throughout childhood (until about 17 years of age), these networks mature by physiologic neurodevelopmental processes, such as selective elimination of neuronal connections and myelination.<sup>13,37–40</sup> As implicated by the interconnecting brain networks between the basal ganglia and cerebellum,<sup>10–12</sup> we also recognized a striking similarity between the presently reported BFMMS age-relatedness and the previously reported age-relatedness of the Scale for the Assessment and Rating of Ataxia (SARA).<sup>21</sup> Comparing age-relatedness between BFMDRS and SARA scales showed that the BFMMS age-relatedness lingered for a longer time course than that of BFMDS and SARA ( $\geq 16$  years versus 12 and 10 years<sup>21</sup> of age, respectively). This could theoretically be attributed to the more detailed subdivision of the BFMMS compared to that of BFMDS and SARA (120 versus 30 and 40 units, respectively).

We recognize several limitations to this study. In studies with (presumably) healthy control children, one could never provide a 100% proof that the included children are really healthy. However, before entering the study, all children fulfilled the predefined inclusion criteria and 2 years after inclusion we checked whether the included children still did. In this 2-year interval after inclusion, two children had developed a neurological diagnosis, consisting of migraine ( $n = 1$ ) and a hernia nuclei pulposi leading to a radicular syndrome ( $n = 1$ ). We checked whether retrospective exclusion of these two children would have changed the outcomes, which was not the case. Furthermore, the pediatric neurologists had also independently provided phenotypic assessments, indicating suspicion of a potential dystonic movement disorder in one child (indicated by two of three pediatric neurologists and subsequently confirmed by M.A.J.T.). The parents of this child had reported no medical complaints. Two years after inclusion, we subsequently checked for the emergence of a dystonic movement disorder by repeating BFMMS scores.

In this child, all “dystonic” features had disappeared. As this child still fulfilled the inclusion criteria, it is tempting to speculate that the transiently observed dystonic features are developmental in origin. Another potential weakness of the study is that the included children exhibited above average educational attainment and that they also participated more frequently in sporting activities than the average Dutch population. However, because we did not observe a correlation between these factors and BFMDRS scores, we would not expect that this has influenced the results. Finally, all assessors had access to the study protocol, implicating that they were aware that they were scoring presumably healthy children. However, we checked the outcomes of all assessors by determining interobserver agreement, both for the total group and also for each assessor subgroup. Furthermore, we obtained a similar age-relatedness from the functional scores that were provided by parents (who were not aware of age-related motor score outcomes). In this perspective, we would therefore suggest that the presented BFMMS age-relatedness can be regarded as indicative.

In conclusion, pediatric BFMDRS outcomes are influenced by age. For optimal interpretation of longitudinal BFMDRS scores in young dystonic children, consideration of pediatric BFMDRS age-relatedness appears advisory. In young children, we hope that further insight in the BFMDRS construct may contribute to adequate and uniform interpretation of longitudinal BFMDRS scores and may facilitate unanimous data entry in international dystonia databases, from pediatric to adult age.

## Author Roles

1. Research Project: A. Conception, B. Organization, C. Execution; 2. Statistical Analysis: A. Design, B. Execution, C. Review and Critique; 3. Manuscript Preparation: A. Writing the First Draft, B. Review and Critique; 4. Video Assessment.

M.J.K.: 1B, 1C, 2A, 2B, 2C, 3A, 4

L.V.: 1B, 1C, 2A, 2B, 2C, 3A, 4

R.B.: 1B, 2A, 2C, 3B, 4

R.J.L.: 3B, 4

H.B.: 2C, 3B

H.E.: 3B, 4

K.J.P.: 3B, 4

M.F.C.: 3B, 4

J.D.S.: 3B, 4

M.A.J.T.: 1A, 3B

D.A.S.: 1A, 1B, 1C, 3A, 3B, 4

## Disclosures

**Funding Sources and Conflict of Interest:** K.J. Peall has received the clinical starter grant from the Academy of Medical Sciences. M.A.J. Tijssen received financial support from Neuro-SIPE NWO Medium Fund, Nuts-Ohra Fund, Prinses Beatrix Fund, Gossweiler Foundation, Science Fund of the Dystonia Society. The authors report no conflicts of interest.

**Financial Disclosures for the previous 12 months:** H. Burger received an honorarium as a statistical consultant for Plos

One. M.F. Contarino is on the advisory boards of Medtronic and Boston Scientific and received speaking fees from Abbvie, Medtronic, Boston Scientific, and ECMT. J.D. Speelman advised Medtronic and Ipsen Pharmaceuticals. M.A.J. Tijssen received fees from Ipsen and Allergan Pharmaceuticals and from Medtronic and Merz. The remaining authors report no financial disclosures.

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## Supporting Information

A video accompanying this article is available in the supporting information here:

**Table S1.** Population Characteristics. Sporting activities are indicated in h/week; school performances are indicated as mean achievements. Dutch population numbers were determined from Trimbos Institute,<sup>32</sup> Central Statistical Office of the Netherlands, and National Kompas.<sup>33</sup> SD, standard deviation; pop., population.

**Table S2.** Video Recording Examination Protocol. \*Speech interview: standard sentences and questions, including questions about problems and the ability to swallow. F, frontal view; P, profile view.

**Table S3.** Multivariable Regression Analysis for the Prediction of BFMMS and BFMD S Scores. Regression analysis results for the effects of the variables on BFMMS and BFMD S scores; when the variables significantly influenced the model (F change), we calculated B (unstandardized coefficient with standard error in parenthesis) and  $\beta$  (standardized regression coefficient). \* $P < 0.001$ .

**Table S4.** Intraclass Correlation Coefficients (ICCs) for BFMMS. Inter- and intraobserver agreement and test–retest reliability for BFMMS (ICCs); latent time interval for determining intraobserver agreement and test–retest reliability was more than 3 weeks. \* $P < 0.05$ . ns, not significant; =, agreement is 1.0 (total agreement).

**Video S1.** Immature motor performances: “drawing a spiral” by a healthy child, 4 years of age.

**Video S2.** Immature motor performances: “finger tapping” by a healthy child, 6 years of age.

**Video S3.** Immature motor performances: “walking” by a healthy child, 8 years of age.

**Video S4.** Developed motor performances: “finger tapping” and “drawing a spiral” by a healthy child, 16 years of age.