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Enhancing prescribing of guideline-recommended medications for ischaemic heart diseases: a systematic review and meta-analysis of interventions targeted at healthcare professionals

Thang Nguyen,1,2 Hoa Q Nguyen,3 Niken N Widyakusuma,4 Thao H Nguyen,3 Tam T Pham,5 Katja Taxis2

ABSTRACT

Objectives Ischaemic heart diseases (IHDs) are a leading cause of death worldwide. Although prescribing according to guidelines improves health outcomes, it remains suboptimal. We determined whether interventions targeted at healthcare professionals are effective to enhance prescribing and health outcomes in patients with IHDs.

Methods We systematically searched PubMed and EMBASE for studies published between 1 January 2000 and 31 August 2017. We included original studies of interventions targeted at healthcare professionals to enhance prescribing guideline-recommended medications for IHDs. We only included randomised controlled trials (RCTs). Main outcomes were the proportion of eligible patients receiving guideline-recommended medications, the proportion of patients achieving target blood pressure and target low-density lipoprotein-cholesterol (LDL-C)/cholesterol level and mortality rate. Meta-analyses were performed using the inverse-variance method and the random effects model. The quality of evidence was assessed using the Grading of Recommendations, Assessment, Development, and Evaluation approach.

Results We included 13 studies, 4 RCTs (1869 patients) and 9 cluster RCTs (15 224 patients). 11 out of 13 studies were performed in North America and Europe. Interventions were of organisational or professional nature. The interventions significantly enhanced prescribing of statins/lipid-lowering agents (OR 1.23; 95% CI 1.07 to 1.42, P=0.004), but not other medications (aspirin/antiplatelet agents, beta-blockers, ACE inhibitors/angiotensin II receptor blockers and the composite of medications). There was no significant association between the interventions and improved health outcomes (target LDL-C and mortality) except for target blood pressure (OR 1.46; 95% CI 1.11 to 1.93; P=0.008). The evidence was of moderate or high quality for all outcomes.

Conclusions Organisational and professional interventions improved prescribing of statins/lipid-lowering agents and target blood pressure in patients with IHDs but there was little evidence of change in other outcomes.

Strenghts and limitations of this study

► This is a systematic review and meta-analysis of randomised controlled trials, conducted following the Cochrane Handbook for Systematic Reviews of Interventions and the Preferred Reporting Items for Systematic Reviews and Meta-Analyses statement.

► This review focused on interventions targeted at healthcare professionals to enhance prescribing of individual medications for acute coronary syndrome. Interventions were classified according to the Cochrane Effective Practice and Organization of Care Review Group. But more detailed analyses, for example, on duration or intensity of intervention implementation, were impossible due to the limited number of studies.

► We may have missed relevant unpublished or locally published studies as we restricted our search to English publications and did not search for grey literature.

INTRODUCTION

Ischaemic heart diseases (IHDs) are a leading cause of death worldwide accounting for 13.2% of all deaths globally.1 IHDs include angina pectoris and myocardial infarction.2 International guidelines recommend using a combination of an antiplatelet agent, a beta-blocker, an ACE inhibitor or an angiotensin II receptor blocker (ACEI/ARB) and an HMG coenzyme A reductase inhibitor (statin) to treat eligible patients with IHDs.3-8 This combination is an effective secondary prevention after myocardial infarction, reducing morbidity and mortality.9-13 Despite such evidence, rates of patients being prescribed medications according to guidelines varied from <5.0% to >95.0%,...
leaving a substantial proportion of patients with IHDs not receiving guideline-recommended care. Changing clinicians’ behaviour to improve prescribing guideline-recommended medications is challenging. Different types of interventions have been developed and classified as professional interventions (eg, education, reminders, audit and feedback), organisational interventions (eg, computerised clinical guidelines, pharmacist-led intervention), financial interventions (eg, financial incentives) and regulatory interventions (eg, cap and copayment policies).

Interventions to improve prescribing guideline-recommended medications for cardiovascular diseases, in general, have been reviewed recently. Moreover, Murphy et al have evaluated the effect of organisational interventions for patients with IHDs. The interventions aimed to improve mortality and hospital admissions and targeted physicians and patients to adhere to recommendations of secondary prevention of IHDs (lifestyle modification, prescribing medications or both). No work has been done synthesising the evidence on interventions to enhance prescribing according to guidelines for patients with IHDs as far as we are aware. In this review, we focus on interventions targeted at health professionals. Other factors influencing prescribing, such as patient behaviour, organisational factors or resource constraints are outside the scope of this review. We conducted a systematic review and meta-analysis to determine whether interventions targeted at healthcare professionals are effective to enhance prescribing and health outcomes in patients with IHDs.

METHODS
We conducted a systematic review and meta-analysis in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) and the Cochrane Handbook for Systematic Reviews of Interventions. We registered our protocol with the International Prospective Register of Systematic Reviews Registry (CRD42016039188). We searched the electronic bibliographic databases PubMed and EMBASE as these are considered to be the most important sources for reports of trials. The search strategy included MeSH terms and relevant keywords in various combinations relating to guidelines, guideline adherence, drug therapy, IHDs and randomised trials (see online supplementary appendix A). We restricted our search to studies carried out in humans and published in English. Studies published between 1 January 2000 and 31 August 2017 were sought. References of included articles were manually screened to identify additional eligible studies.

We included original studies reporting results of randomised controlled trials (RCTs) or cluster randomised controlled trials (cluster RCTs) in patients with IHDs eligible for receiving secondary preventive treatment. Studies had to evaluate interventions targeted at healthcare professionals to enhance prescribing of guideline-recommended medications. The trials had to include at least one prospectively assigned control group. The control group had to receive usual care (not receiving the intervention), or an intervention of lower intensity or shorter duration than the intervention group. Studies had to report patient-level outcomes. We excluded duplicate reports, post hoc analyses or abstracts from meeting proceedings unless published as full-text reports in a peer-reviewed journal. We excluded studies on patients receiving acute treatment in hospital only, or interventions predominantly targeting patient medication-taking behaviour or lifestyle modifications.

All titles and abstracts retrieved from the electronic searches were archived in the web-based bibliography and database manager RefWorks. After removing duplicates, two reviewers (TN and HQN) independently screened the titles and abstracts. They also independently assessed the full text of potentially eligible studies. Disagreements between the reviewers whether to include or exclude a study were resolved by consensus.

Two reviewers (TN and NNW) independently extracted data from the trials’ primary texts, the online supplementary appendices and protocols using a data abstraction form. We extracted the following information: trial name, year of publication, sources of funding, setting and time of recruitment, study design, study population characteristics, details of the intervention and control conditions, main outcomes and evidence for assessment of the risk of bias. Disagreements were resolved by discussion with a third reviewer (KT).

Two reviewers (TN and NNW) independently assessed the risk of bias of each study using the tool of the Cochrane Effective Practice and Organization of Care Review Group (EPOC). The nine standard criteria were: (1) random sequence generation, (2) allocation sequence concealment, (3) similarity of baseline outcome measures, (4) similarity of baseline characteristics, (5) blinding of outcome assessment, (6) adequately addressing incomplete outcome data, (7) adequate protection against contamination, (8) free from selective reporting and (9) free from other risks of bias (eg, recruitment bias or not adjusting for clustering effect in cluster RCTs). Disagreements were resolved by discussion with a third reviewer (KT). We judged trials with four or more high-risk domains, or three or more high-risk domains plus three or more unknown domains as having a high risk of bias.

The primary outcomes were the proportion of eligible patients receiving the following guideline-recommended medications: aspirin/antiplatelet agents, beta-blockers, ACEI/ARBs, statins/lipid-lowering agents and a composite of these medications. The secondary outcomes were: the proportion of eligible patients achieving target blood pressure and target LDL-C/cholesterol level, and the mortality rate.

The interventions were classified according to the taxonomy of the EPOC as professional, financial,
organisational or regulatory interventions. We performed meta-analyses for outcomes when the necessary data were available. Meta-analyses were performed in the Review Manager V.5.3 (RevMan 5) using the inverse-variance method and the random effects model. The OR with corresponding 95% CI was calculated for each outcome of interest to generate a forest plot. For studies with more than two trial groups, we combined relevant groups to create a single pair-wise comparison. A Z-test was used to assess the statistical significance of the results of the meta-analysis with a two-tailed P value of <0.05. The intraclass correlation coefficients (ICCs) for cluster RCTs were used to calculate the effective sample size to ensure the clustering effect was taken into account in our analyses. When an ICC was not reported in a cluster RCT, we contacted the trial authors. In case of non-response, we used the mean of corresponding ICCs reported in the other included cluster RCTs to adjust for the clustering effect.

Two reviewers (KT and TN) independently assessed the quality of evidence across included studies of all outcomes of interest using the Grading of Recommendation, Assessment, Development, and Evaluation (GRADE) approach. The following criteria were used: serious limitations in study design and implementation, indirectness, substantial heterogeneity, imprecision and publication bias. The GRADE approach specifies four levels of quality: high, moderate, low and very low. The quality rating was downgraded by one level for each factor having a serious limitation, up to a maximum of three levels for all factors. Heterogeneity across trials for each outcome of interest was investigated using the Cochran’s Q test and was measured by the I^2 statistic. An I^2 exceeding 50% indicated substantial statistical heterogeneity. Publication bias was evaluated visually by inspecting funnel plots and quantified by the Egger’s test for outcomes comprising at least 10 trials.

We performed subgroup analyses and sensitivity analyses when the necessary data were available. Subgroup analyses were performed for type of study designs, type of intervention, comparators and setting of the intervention. We examined the robustness of our findings in sensitivity analyses excluding studies with high overall risk of bias, and analyses without adjusting for clustering effect.

RESULTS

The search of PubMed and EMBASE databases provided a total of 8424 citations, and 452 citations were added from the lists of references from included studies. After removing duplicates, 7555 remained. Of those, 7219 papers were discarded after screening titles and abstracts. The full text of 316 studies was examined in more detail, 303 studies did not meet the inclusion criteria. A total of 13 studies were identified for inclusion in the review (figure 1). These were 4 RCTs involving 1869 patients and 9 cluster RCTs involving 599 patients and 9 cluster RCTs43 47 50 52 53 55–58 involving 599 patients and 9 cluster RCTs43 47 50 52 53 55–58 involving 599 healthcare centres and 15224 patients. Trials were carried out between 1997 and 2012 and published between 2001 and 2015. Control groups received usual care (nine studies43 45 49–52 55 58 59) or less intensive interventions (four studies47 53 56 57). Seven studies43 49 52 53 55 57 59 reported patients’ health outcomes (table 1). The overall risk of bias was rated as low in all included studies (table 1 and more details in online supplementary appendix B).

Five studies45 49 50 52 59 used organisational interventions, four studies43 51 53 55 58 professional interventions and four studies43 47 56 57 a combination of organisational and professional interventions. Distribution of educational materials, educational outreach visits, audit and feedback and reminders were the four professional interventions most frequently used. Continuity of care, communication and case discussions between distant healthcare professionals were the two organisational interventions most frequently used (table 2 and more details in online supplementary appendix C).

Interventions had no significant effect on prescribing guideline-recommended medications, that is, there was no significant difference in the proportion of eligible patients receiving guideline-recommended medications between intervention and control groups except for statins/lipid-lowering agents. The findings were aspirin/antiplatelet agents (OR 1.13; 95% CI 0.87 to 1.47; P=0.360), beta-blockers (OR 1.13; 95% CI 0.93 to 1.37; P=0.230), ACEIs/ARBs (OR 1.04; 95% CI 0.88 to 1.23; P=0.620) and statins/lipid-lowering agents (OR 1.23; 95% CI 1.07 to 1.42; P=0.004), the composite of medications (OR 1.07; 95% CI 0.73 to 1.58; P=0.720). The evidence was of moderate or high quality for the primary outcomes (figure 2 and table 3).
Table 1  Characteristics of included studies

<table>
<thead>
<tr>
<th>No.</th>
<th>Source</th>
<th>Study design</th>
<th>Study period</th>
<th>Country Setting of recruitment</th>
<th>Diagnosis</th>
<th>Intervention group</th>
<th>Control group</th>
<th>Primary outcome</th>
<th>Secondary outcome</th>
<th>Overall risk of bias</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>Berger et al.</td>
<td>Cluster RCT</td>
<td>2011–2012</td>
<td>Brazil Hospital, Spain</td>
<td>ACS, OI plus PI</td>
<td>602 (17) 62 (13)</td>
<td>68.6 UC 548 (17) 62 (13)</td>
<td>68.6 ASA, BB, ACEI, statin</td>
<td>Low 35-day mortality</td>
<td>No Low</td>
</tr>
<tr>
<td>2</td>
<td>Bond et al.</td>
<td>RCT</td>
<td>2002–2004</td>
<td>UK GP/PCP HD</td>
<td>CI 941 68.7 (9.2)</td>
<td>67.4 UC 500 68.9 (9.1)</td>
<td>70.6 ASA, BB, CLO, BB, ACEI, statin</td>
<td>Low 4 month mortality</td>
<td>No Low</td>
<td></td>
</tr>
<tr>
<td>3</td>
<td>Flisser et al.</td>
<td>Cluster RCT</td>
<td>2008</td>
<td>NA France, Italy, Spain</td>
<td>NSTEACS OI plus PI</td>
<td>722 (19) 65.6 (10.9)</td>
<td>67.2 UC 479 (19) 66.1 (10.6)</td>
<td>72.2 ASA, BB, ACEI, statin</td>
<td>Low 30-day mortality</td>
<td>No Low</td>
</tr>
<tr>
<td>4</td>
<td>Gaglia et al.</td>
<td>RCT</td>
<td>2009–2010</td>
<td>Norway Hospital</td>
<td>IHD OI 48</td>
<td>63.9 (9)</td>
<td>72 UC 46 63.4 (9.9)</td>
<td>72 ASA, BB, Target LDL-C</td>
<td>Low 30-day mortality</td>
<td>No Low</td>
</tr>
<tr>
<td>5</td>
<td>Flaherty et al.</td>
<td>Cluster RCT</td>
<td>2001–2003</td>
<td>12 UK GP/PCP HD</td>
<td>MI 232 (184) 68.3 (11.3) 66 4</td>
<td>227 (210) 67.3 (12.1) 63.9 ASA, BB, ACEI, statin</td>
<td>Low 30-day mortality</td>
<td>No Low</td>
<td></td>
<td></td>
</tr>
<tr>
<td>6</td>
<td>Hung et al.</td>
<td>RCT</td>
<td>2004–2007</td>
<td>Taiwan Hospital</td>
<td>IHD Pl</td>
<td>92 67 (10) 71.7</td>
<td>71 UC 102 66.1 (11)</td>
<td>75.2 LL</td>
<td>Low 30-day mortality</td>
<td>No Low</td>
</tr>
<tr>
<td>7</td>
<td>Khunti et al.</td>
<td>Cluster RCT</td>
<td>2002–2008</td>
<td>USA GP/PCP HD</td>
<td>MI 3060 (84)</td>
<td>98.8 UC 2911 (84) 98.7</td>
<td>82.1 APA, BB, ACEI, ARB, statin</td>
<td>Low 30-target BP, LDL-C</td>
<td>Low 30-day mortality</td>
<td>No Low</td>
</tr>
<tr>
<td>8</td>
<td>Levine et al.</td>
<td>Cluster RCT</td>
<td>2005–2008</td>
<td>Canada Hospital</td>
<td>MI 682 (7) 66.4 (5.9)</td>
<td>67 UC 559 (7) 66.5 (5.4)</td>
<td>67 APA, LL</td>
<td>Low 30-day mortality</td>
<td>No Low</td>
<td></td>
</tr>
<tr>
<td>9</td>
<td>McAlister et al.</td>
<td>Cluster RCT</td>
<td>2005–2008</td>
<td>Canada Hospital</td>
<td>MI 1422 (10) NR NR</td>
<td>98.8 UC 1166 (10) NR NR</td>
<td>98.7 APA, BB, ACEI, ARB, statin</td>
<td>Low 30-target BP, LDL-C</td>
<td>Low 30-day mortality</td>
<td>No Low</td>
</tr>
<tr>
<td>10</td>
<td>Moher et al.</td>
<td>Cluster RCT</td>
<td>2002–2003</td>
<td>Denmark Hospital</td>
<td>ACS</td>
<td>665 (7) 65.6 (8.5) 71</td>
<td>71 UC 157 (9) 64.9 (8.4)</td>
<td>82.1 APA, BB, ACEI, ARB, statin</td>
<td>Low 30-target BP, LDL-C</td>
<td>Low 30-day mortality</td>
</tr>
<tr>
<td>11</td>
<td>Ornstein et al.</td>
<td>Cluster RCT</td>
<td>2000–2002</td>
<td>USA Hospital</td>
<td>ACS PI 1422 (10) NR NR</td>
<td>98.8 UC 1166 (10) NR NR</td>
<td>98.7 APA, BB, ACEI, ARB, statin</td>
<td>Low 30-target BP, LDL-C</td>
<td>Low 30-day mortality</td>
<td>No Low</td>
</tr>
<tr>
<td>12</td>
<td>Sondergaard et al.</td>
<td>Cluster RCT</td>
<td>2000–2003</td>
<td>NA</td>
<td>ACS</td>
<td>665 (7) 65.6 (8.5) 71</td>
<td>71 UC 157 (9) 64.9 (8.4)</td>
<td>82.1 APA, BB, ACEI, ARB, statin</td>
<td>Low 30-target BP, LDL-C</td>
<td>Low 30-day mortality</td>
</tr>
<tr>
<td>13</td>
<td>Yorio et al.</td>
<td>RCT</td>
<td>2003–2004</td>
<td>USA Hospital</td>
<td>ACS</td>
<td>55.9 (11.3) 66.7</td>
<td>68 UC 162 (14) 56.2 (10.8) 57.3 APA, LL</td>
<td>Low 30-target BP, LDL-C</td>
<td>Low 30-day mortality</td>
<td>No Low</td>
</tr>
</tbody>
</table>

ACEI, ACE inhibitors; ACS, acute coronary syndrome; APA, antiplatelet agents; ARB, angiotensin II receptor blockers; ASA, aspirin; BB, beta-blockers; BP, blood pressure; CLO, clopidogrel; GP, general practice; HD, high-density lipoprotein cholesterol; IQR, interquartile range; LDL-C, low-density lipoprotein cholesterol; MI, myocardial infarction; NA, not applicable; NR, not reported; NSTEACS, non-ST-elevation acute coronary syndrome; OI, organisational intervention; PCP, primary care practice; PI, professional intervention; RCT, randomised controlled trials; SBP, systolic blood pressure; UC, usual care.
<table>
<thead>
<tr>
<th>No.</th>
<th>Source</th>
<th>Setting of intervention implementation</th>
<th>Intervention carried out by</th>
<th>Distribution of educational materials</th>
<th>Educational meeting</th>
<th>Educational outreach visits</th>
<th>Local opinion leaders</th>
<th>Audit and feedback</th>
<th>Reminders</th>
<th>Revision of professional roles</th>
<th>Clinical multidisciplinary teams</th>
<th>Continuity of care</th>
<th>Communication and case discussion between distant healthcare professionals</th>
<th>Presence and organisation of quality monitoring mechanisms</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>Berwanger et al</td>
<td>Hospital</td>
<td>Nurse and physician</td>
<td>x</td>
<td>x</td>
<td>x</td>
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<tr>
<td>2</td>
<td>Bond et al</td>
<td>Pharmacy</td>
<td>Community pharmacist</td>
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<td>x</td>
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<td>x</td>
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<tr>
<td>3</td>
<td>Flather et al</td>
<td>Hospital</td>
<td>Cardiologist, nurse and manager</td>
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<td>4</td>
<td>Garcia et al</td>
<td>GP/PCP</td>
<td>Hospital pharmacist</td>
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<td>GP/PCP</td>
<td>Cardiologist</td>
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<tr>
<td>6</td>
<td>Hung et al</td>
<td>Hospital</td>
<td>Reminder system</td>
<td></td>
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<td>x</td>
</tr>
<tr>
<td>7</td>
<td>Khunti et al</td>
<td>GP/PCP</td>
<td>Nurse</td>
<td></td>
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<td>x</td>
</tr>
<tr>
<td>8</td>
<td>Levine et al</td>
<td>GP/PCP</td>
<td>Internet-delivered intervention system</td>
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<tr>
<td>9</td>
<td>McAlister et al</td>
<td>GP/PCP</td>
<td>Leader</td>
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</tr>
<tr>
<td>10</td>
<td>Moher et al</td>
<td>GP/PCP</td>
<td>General practitioner and nurse</td>
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<td>11</td>
<td>Ornstein et al</td>
<td>GP/PCP</td>
<td>Not specified</td>
<td>x</td>
<td>x</td>
<td>x</td>
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<td></td>
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<td></td>
<td></td>
<td>x</td>
</tr>
<tr>
<td>12</td>
<td>Sondergaard et al</td>
<td>GP/PCP</td>
<td>Not specified</td>
<td>x</td>
<td>x</td>
<td>x</td>
<td></td>
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</tr>
<tr>
<td>13</td>
<td>Yorio et al</td>
<td>Cardiology clinic</td>
<td>Nurse or clinical pharmacist</td>
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</tbody>
</table>

*The interventions were classified according to the taxonomy of the Cochrane Effective Practice and Organization of Care Review Group. GP, general practice; PCP, primary care practice.
Figure 2  Primary outcomes of intervention vs control. ACEI, ACE inhibitors; ARB, angiotensin II receptor blocker.
Table 3 Summary of findings and quality assessment

<table>
<thead>
<tr>
<th>Outcome</th>
<th>Study design</th>
<th>Indirectness</th>
<th>Substantial statistical heterogeneity</th>
<th>Imprecision</th>
<th>Publication bias</th>
<th>Summary of findings</th>
<th>Quality of the evidence (GRADE)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
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<td></td>
<td>Illustrative comparative risks (95% CI)</td>
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<td></td>
<td></td>
<td></td>
<td>Assumed risk in comparison</td>
<td>Corresponding risk in intervention</td>
</tr>
<tr>
<td>ASA/APA</td>
<td>Yes*</td>
<td>No</td>
<td>No</td>
<td>No</td>
<td>No</td>
<td>851 per 1000</td>
<td>866 per 1000 (832 to 894)</td>
</tr>
<tr>
<td>BB</td>
<td>Yes*</td>
<td>No</td>
<td>No</td>
<td>No</td>
<td>No</td>
<td>840 per 1000</td>
<td>856 per 1000 (830 to 878)</td>
</tr>
<tr>
<td>ACEI/ARB</td>
<td>No</td>
<td>No</td>
<td>No</td>
<td>No</td>
<td>Unknown†</td>
<td>735 per 1000</td>
<td>743 per 1000 (709 to 773)</td>
</tr>
<tr>
<td>Statin/LLA</td>
<td>Yes*</td>
<td>No</td>
<td>No</td>
<td>No</td>
<td>No</td>
<td>770 per 1000†</td>
<td>805 per 1000 (782 to 826)</td>
</tr>
<tr>
<td>Composite</td>
<td>No</td>
<td>No</td>
<td>No</td>
<td>No</td>
<td>Unknown†</td>
<td>566 per 1000</td>
<td>583 per 1000 (488 to 673)</td>
</tr>
<tr>
<td>Target BP</td>
<td>Yes*</td>
<td>No</td>
<td>No</td>
<td>No</td>
<td>Unknown†</td>
<td>432 per 1000</td>
<td>526 per 1000 (458 to 595)</td>
</tr>
<tr>
<td>Target LDL-C/cholesterol</td>
<td>Yes*</td>
<td>No</td>
<td>No</td>
<td>No</td>
<td>Unknown†</td>
<td>704 per 1000†</td>
<td>714 per 1000 (682 to 744)</td>
</tr>
<tr>
<td>Mortality</td>
<td>No</td>
<td>No</td>
<td>No</td>
<td>Yes§</td>
<td>Unknown†</td>
<td>84 per 1000</td>
<td>67 per 1000 (42 to 104)</td>
</tr>
</tbody>
</table>

Patient or population: patients with ischaemic heart diseases.
Comparison: usual care or less intensive intervention.
Intervention: interventions intended to improve prescribing guideline-recommended medications and patients’ health outcomes.
§Included study had few events and wide CI.
*More than one-third of studies had recruitment bias.
†Did not perform Egger’s test because of number of studies <10.
‡Did not include the study by Hung et al because its population was the patients not receiving statin/LLA appropriately at baseline.
Setting: hospitals, general practices/primary care practices, cardiology clinics or pharmacies.
ACEI, ACE inhibitors; APA, antiplatelet agents; ARB, angiotensin II receptor blockers; ASA, aspirin; BB, beta-blockers; BP, blood pressure; GRADE, grading of recommendations assessment, development and evaluation; LDL-C, low-density lipoprotein cholesterol; LLA, lipid-lowering agents.
The interventions significantly increased the proportion of patients achieving target blood pressure (OR 1.46; 95% CI 1.11 to 1.93; \( P = 0.008 \)), but there was no significant difference in the proportion of patients achieving target LDL-C/cholesterol (OR 1.05; 95% CI 0.90 to 1.22; \( P = 0.550 \)), and in mortality rate (OR 0.78; 95% CI 0.48 to 1.27; \( P = 0.320 \)) between intervention and control groups. The evidence was of moderate quality for the secondary outcomes (figure 3 and table 3).

No substantial statistical heterogeneity was detected in our study outcomes (all eight \( I^2 \) values were <50%) (figure 2 and figure 3). The publication bias was rated as no risk (in aspirin/antiplatelet agents, beta-blockers and statins/lipid-lowering agents) and unknown risk (in the other outcomes) (see online supplementary appendix D). In subgroup analyses, there was no significant difference in the effect of the interventions on prescribing guideline-recommended medications and patients’ health outcomes between subgroups with all \( P \) values for the interaction of >0.05. No subgroup analysis could be done for the composite of medications and mortality rate as there were only two studies available for each of these outcomes (see online supplementary appendix E). We did not perform sensitivity analyses excluding studies with high overall risk of bias because all included studies were rated as low risk. The findings of all outcomes did not change in sensitivity analyses when not adjusting for clustering effects (see online supplementary appendix F).

**DISCUSSION**

Interventions to enhance prescribing guideline-recommended medications for patients with IHDs were of organisational or professional nature. The interventions...
For example, the benefits of more intensive therapy with not adequate dosing had been achieved was not measured agents, but target blood pressure improved. Whether or not on LDL-C/cholesterol level. In contrast, interventions impact on prescribing of statins/lipid-lowering agents, but not other medications. There was no significant difference between subgroups of interventions (professional, organisational and professional plus organisational). Subgroup analyses showed that there was no significant difference between subgroups of interventions (professional, organisational and professional plus organisational). But more detailed analyses, for example, on duration or intensity of the intervention, were impossible due to the limited number of studies. The length of patient follow-up varied across studies. This issue might increase the clinical heterogeneity of outcomes measured. Fourth, we included studies reporting patient-level outcomes, and excluded studies only reporting cluster-level outcomes (eg, hospital and practice performance scores). Fifth, we performed multiple statistical tests which increased the risk of type I error. Adjustment for multiple testing is debatable. In our study, three out of four primary outcomes were not significant, P value threshold adjustment would be too conservative. Finally, our review included only studies published in English and we did not search for grey literature. So we

Our findings are consistent with previous systematic reviews28 29 reporting professional and organisational as the two main types of interventions to improve healthcare professionals' adherence to cardiovascular disease guidelines. Our study and a systematic review by Jeffery et al38 showed only some significant improvements. A systematic review by Unverzagt et al28 in contrast, showed that a provider reminder system, audit and feedback, provider education or organisational change were effective interventions. However, results are difficult to compare as we measured different outcomes. We analysed the improvement of prescribing for each medication separately while both review articles28 29 took all medication together. Moreover, we focused on patients with IHDs, whereas previous reviews28 29 included different cardiovascular diseases. Although programmes promoting guidelines such as the Guidelines Applied in Practice and Get With The Guidelines programmes also involving organisational and professional interventions demonstrated that it was possible to improve quality of care,67 the design of RCT is needed to confirm the improvement.

Several issues need to be addressed in our study. First, there were seven studies rated as having a high risk of other bias. Of these studies, six cluster RCTs30 32 33 34 36 38 had a high risk of recruitment bias. In those studies, patients were recruited after the clusters had been randomised and therefore, the knowledge of whether a cluster belonged to the intervention or control group could have affected patient recruitment. Farrin et al68 showed this in a trial of low back pain randomised by primary care practice; a greater number of less severe participants were recruited to the ‘active management’ practices. However, we did not find significant differences in outcomes between RCTs,45 49 51 59 and cluster RCTs.45 47 50 52 53 55–58 Second, there were some cluster RCTs,47 50 53 56 58 which did not report the ICCs. We used the mean ICCs for corresponding outcomes reported in the other included studies.39 The sensitivity analyses without adjusting for clustering effects showed similar results. The heterogeneity became substantial for the outcomes of aspirin/antiplatelet agents, the composite of medications and target LDL-C/cholesterol. But overall, the sensitivity analyses confirmed the robustness of our findings. Third, we included studies of all types of interventions targeted at healthcare professionals in the meta-analyses. Subgroup analyses showed that there was no significant difference between subgroups of interventions (professional, organisational and professional plus organisational). But more detailed analyses, for example, on duration or intensity of the intervention, were impossible due to the limited number of studies. The length of patient follow-up varied across studies. This issue might increase the clinical heterogeneity of outcomes measured. Fourth, we included studies reporting patient-level outcomes, and excluded studies only reporting cluster-level outcomes (eg, hospital and practice performance scores).9 70 Fifth, we performed multiple statistical tests which increased the risk of type I error. Adjustment for multiple testing is debatable. In our study, three out of four primary outcomes were not significant, P value threshold adjustment would be too conservative. Finally, our review included only studies published in English and we did not search for grey literature. So we
may have missed relevant unpublished or locally published studies.

Our results have several implications for practice and research. Eleven out of 13 studies come from North America and Europe, which limits the generalisability of our results to the rest of the world. There may be a need to develop new interventions, especially for low-income and middle-income countries which have a rising burden of ischaemic heart diseases. There are some types of interventions such as financial and regulatory that have not been tested in this group. Selecting an intervention to enhance prescribing according to guidelines should be based on the local context. Interventions need to consider a range of barriers to change prescribing, including barriers related to patients, organisation of the healthcare system and resource constraints. Finally, improving guideline adherence may include strategies for improving clinicians’ awareness, agreement and adoption of guidelines. The cost-effectiveness of such interventions should also be evaluated.

CONCLUSIONS

In conclusion, a number of organisational and professional interventions improved prescribing of statins/lipid-lowering agents and target blood pressure in patients with IHDs, but there was little evidence of change in other outcomes.

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Enhancing prescribing of guideline-recommended medications for ischaemic heart diseases: a systematic review and meta-analysis of interventions targeted at healthcare professionals

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