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**HARALD J. VAN HEERDE, MAARTEN J. GIJSENBERG, MARNIK G. DEKIMPE, and JAN-BENEDICT E.M. STEENKAMP**

Firms are under increasing pressure to justify their marketing expenditures. This evolution toward greater accountability is reinforced in harsh economic times when marketing budgets are among the first to be reconsidered. To make such decisions, managers must know whether, and to what extent, marketing's effectiveness varies with the economic tide; however, surprisingly little research addresses this issue. Therefore, the authors conduct a systematic investigation of the business cycle's impact on the effectiveness of two important marketing instruments: price and advertising. To do so, they estimate time-varying short- and long-term advertising and price elasticities for 150 brands across 36 consumer packaged goods categories, using 18 years of monthly U.K. data from 1993 to 2010. The long-term price sensitivity tends to decrease during economic expansions, whereas long-term advertising elasticities increase. During contractions, the long-term own and cross price elasticities increase. Moreover, throughout the observation period, the short-term price elasticity became significantly stronger. Finally, patterns differ across categories and brands, which presents opportunities for firms that know how to ride the economic tide.

*Keywords:* marketing-mix effectiveness, market-response models, recession, business cycle.

**Price and Advertising Effectiveness over the Business Cycle**

During economic downturns, managers feel the urge to make changes to their marketing mix. Indeed, in a survey of 1400 managers, 96% indicated that they had made changes in the wake of the global financial crisis (McKinsey & Company 2009). One of the most common actions was the reduction of marketing support for the firm’s brands (Deleersnyder et al. 2009; Srinivasan, Rangaswamy, and Lilien 2005). According to MagnaGlobal Interpublic, in 2009, worldwide advertising expenditures dropped 10.8% (in the United States, –15.8%) (Advertising Age 2011). In their 2003 article, Marn, Roegner, and Zawada discuss how many managers increase their prices during a contraction to offset the revenue losses caused by reduced sales volumes. Yet, recent academic studies have shown that these responses to the business cycle may be counterproductive (Deleersnyder et al. 2009; Lamey et al. 2012; Steenkamp and Fang 2011). This leaves brand managers struggling with questions whether, in what direction, and how much economic conditions influence the effectiveness of key marketing instruments such as advertising and price. This is an unfortunate state of affairs, because knowledge of marketing’s impact is crucial to make informed decisions (Hanssens 2009).

The purpose of this study is to provide brand managers with these insights. More specifically, we investigate whether and how short- and long-term advertising and price

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elastici- ties change over the business cycle. We conduct a
large-scale empirical study spanning 18 years (1993–2010)
of monthly data—including the effects of the global finan-
cial crisis—for 150 brands in the United Kingdom, across
36 consumer packaged goods (CPG) categories. The single-
source data enable us to use a consistent measurement and
modeling scheme across all brands, which eliminates the
confounding factors identified by earlier meta-analyses on
price and advertising (Bijmolt, Van Heerde, and Pieters
2005; Sethuraman, Tellis, and Briesch 2011).

PREVIOUS RESEARCH ON BUSINESS CYCLES

In recent years, marketing researchers have paid increas-
ing attention to the link between marketing phenomena and
macroeconomic fluctuations. Deleersnyder et al. (2004)
study the evolution in consumer durables sales during
expansions and contractions, whereas Lamey et al. (2007)
document the effect of the business cycle on the evolution
of private-label share in several Western countries. Deleers-
snyder et al. (2009) consider cross-national variation in the
cyclical fluctuations in aggregate (country-level) advertis-
ing spending, and Millet, Lamey, and Van den Bergh (2012)
show that consumers’ motivational orientations differ across
economic expansion and contraction periods. These studies
document that over-time variation in the level of the focal
variable is strongly related to the business cycle.

A few studies have begun to investigate the implications
of pro- versus countercyclical marketing spending. Franken-
berger and Graham (2003) find that advertising expenditures
during economic downturns offer more financial benefits
(e.g., operating income, shareholder value) than advertising
expenditures in expansions. Deleersnyder et al. (2009)
report that companies whose advertising expenditures
behave procyclically show poorer stock-price performance
than companies that set their advertising investments inde-
pendent of business cycles. Srinivasan, Rangaswamy, and
Lilien (2005) document that firms that rate high on the latent
construct of “proactive marketing response in a recession”
report higher business performance. Finally, Steenkamp and
Fang (2011) find that an increase in share of voice has a
stronger impact on profit and market share than increasing
advertising share in expansions.

Collectively, these studies suggest that the effectiveness
of marketing-mix instruments differs between expansion
and contraction periods. However, the researchers con-
ducted all studies at the firm level. While insights derived
from firm-level analyses are undoubtedly important for the
boardroom, they do not necessarily help individual brand or
category managers with decision making. Especially for
firms that are “a house of brands,” as most CPG companies
are (Kapferer 2008), aggregate firm-level results may be
less relevant because the competitive battle is fought at the
individual brand level. The results might even be misleading,
due to aggregation bias (Christen et al. 1997). Relatedly,
previous studies often focused on aggregate accounting/
financial metrics (stock-price reaction, firm profit). Although
Steenkamp and Fang (2011) also consider firm market
share, this measure is an aggregate across the different
industries (Standard Industrial Classification codes) in
which the company is active.

Moreover, the aforementioned studies focus on advertising. Although advertising is undoubtedly one of the most
important marketing weapons, existing empirical general-
izations indicate that price elasticities are typically 10–20
times larger than advertising elasticities (Bijmolt, Van
Heerde, and Pieters 2005; Sethuraman, Tellis, and Briesch
2011). This supports the argument that pricing decisions are
even more important for brand performance than advertis-
ing decisions.

Gordon, Goldfarb, and Li (2013) analyze how price elastic-
ties vary with the state of the economy using six years of
Information Resources Inc. panel data, finding that price
sensitivity rises when the economy weakens. Their price
elasticity measures the current-period (short-term) effect but
does not address what happens to long-term price elasticity.
In addition, this study does not examine how short- and
long-term advertising elasticity changes with the business
cycle, nor does it allow for differential effects of expansions
versus contractions. Moreover, their data period (2001–
2006) is relatively short in terms of both the number of ups
and downs in the economy and the magnitude of these
changes. Finally, they do not study differences between
brands.

Our study is designed to address these limitations in pre-
vious research. We are the first to document how brand-
level elasticities in both price and advertising evolve over
the business cycle in the short run as well as the long run.
To do so, we employ a data set that is unprecedented in its
composition in that we estimate, using monthly data, elast-
cities for 150 brands over a period of 18 years, covering
multiple expansion and contraction periods.

RESEARCH FRAMEWORK

Our study focuses on the effect of the business cycle on the
short- and long-term effectiveness of own brand adver-
sising and price on brand sales. Previous research has estab-
lished that performance metrics such as category sales
(Deleersnyder et al. 2004) and private-label share (Lamey
et al. 2007, 2012) as well as marketing-support variables
(Deleersnyder et al. 2009) react differently (i.e., asymmetri-
cally) to economic expansions and contractions. This also
applies to consumers’ motivational orientations (Millet,
Lamey, and Van den Bergh 2012). Given these findings, we
do not impose symmetry in effectiveness either and allow
for different response parameters along both phases of the
business cycle.

Although this article’s primary focus is the moderating
effect of the business cycle on a brand’s own price and
advertising effectiveness, we also consider three additional
managerially important issues. First, we investigate whether
the effectiveness of advertising or price systematically
declines or increases with the passage of time. Second, we
examine whether advertising systematically moderates a
brand’s price elasticity. Third, we examine whether the
effect of competitor advertising and price varies across the
business cycle. Figure 1 provides a schematic overview of
the major aspects of our study.

Own Price Effectiveness over the Business Cycle

Own price elasticity. In a large-scale meta-analysis, Tellis
(1988) covers 367 elasticities related to 220 different brands
or markets, showing a grand mean of –1.76. More recently,
from a meta-analysis of 1851 elasticities, Bijmolt, Van
Heerde, and Pieters (2005) report an average price elasticity
of $-2.62$ (SD = 2.21). The latter authors have also shown that over time, consumers have become increasingly price sensitive. This is consistent with Mela, Jedidi, and Bowman (1998), who find that households develop price expectations on the basis of their previous exposure to promotions over a long period of time. They report that offering frequent price promotions (as has become increasingly common in the grocery industry) leads to (1) a reduced likelihood of purchase incidence on a given shopping trip but (2) an increase in the quantity bought when households decide to buy, typically with a promotional offer. In line with these findings, we expect that the magnitude of price elasticity increases over time (i.e., becomes more negative).

Role of the business cycle. We predict that during economic downturns, price elasticity becomes stronger (i.e., more negative). Consumers’ disposable income is lower in such periods, creating a higher level of price awareness and fostering a search for lower prices (Estelami, Lehmann, and Holden 2001). Consumers are looking more often for price deals (Quelch 2008) and switch more frequently to lower-priced private-label offerings (Lamey et al. 2007). We therefore expect an increase in the magnitude of both short- and long-term price elasticity (i.e., more negative value) during economic expansions.

Own Advertising Effectiveness over the Business Cycle

Own advertising elasticity. Allenby and Hanssens (2004) review advertising-response research from the last 25 years and find that short-term advertising elasticities for established products are very small (approximately .01). In a recent meta-analysis, Sethuraman, Tellis, and Briesch (2011) find a mean long-term elasticity of .24 across 402 observations, with 40% of these elasticities between 0 and .1. Srinivasan, Vanhuele, and Pauwels (2010) report an average long-term advertising elasticity of .036 across 74 brands in four CPG categories. Sethuraman, Tellis, and Briesch (2011) also find that advertising elasticity is lower in more recent studies, which suggests that advertising elasticity declines over time. This is “because of increased competition, ad clutter, the advent of the Internet as an alternate information source, and the consumer’s ability to opt out of television commercials through devices such as TiVo” (Sethuraman, Tellis, and Briesch 2011, p. 460).

Role of the business cycle. Whereas price elasticities are expected to become stronger during economic downturns, there are opposing predictions for advertising elasticities. On the one hand, advertising elasticities may increase during a contraction. Decreasing advertising budgets (Deleersnyder et al. 2009) will result in less competitive clutter (Danahe, Bonfrer, and Dhar 2008), which may make it easier for customers to notice individual firms. In addition, media rates are lower during contractions (Parekh 2009), meaning that the same advertising dollar buys more advertising coverage. Both factors contribute to an increased effectiveness of firms’ advertising dollars during a contraction.

On the other hand, Sethuraman, Tellis, and Briesch (2011, p. 46) argue that during contractions, consumers become more price conscious and generally tend to ignore image-based advertising. Moreover, in contractions, brands relying heavily on advertising may be perceived as being less sympathetic to the consumer’s tight economic situation.
select up to five leading national brands. Another requirement is that a brand needs to advertise sufficiently often: at least twice during a contraction and at least twice during an expansion. In total, we included 150 national brands. To illustrate the range of products available in our data set, we have grouped them into broader product classes—food, beverages, household care, and personal care. Table 1 shows the number of categories in these product classes and some illustrative examples for several sample categories. The table highlights that our data cover many well-known household names.

We obtained price information on the 150 brands from Kantar Worldpanel UK and acquired information about their monthly advertising expenditure from Nielsen Media (United Kingdom). All marketing-mix series are inflation-adjusted using the U.K. Consumer Price Index, which we obtained from the Organisation for Economic Co-operation and Development.

We use data on real gross domestic product (GDP) as a proxy for general economic activity. Stock and Watson (1999) have shown the cyclical component of the GDP series to be a good indicator of overall economic cycle. We obtained GDP data, expressed in constant prices, from the Organisation for Economic Co-operation and Development. The time span considered covers various economic conditions ranging from periods of relative stability to periods of economic decline (including the recent global financial crisis) and periods with considerable growth.

**METHODOLOGY**

Given our research objectives, our modeling approach needs to address several issues. First, our model should provide short- and long-term advertising and price elasticities. Second, it should allow for interactions between these elasticities and the state of the economy. Third, because consumers may react differently to economic contractions and expansions, it should allow for asymmetric effects. Fourth, because it is unlikely that expansions and contractions affect all brands in the same way, the response parameters should be able to vary across brands. Fifth, we need to allow for possible error correlations between brands from the same

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1 The maximum number of brands to be considered is $36 \times 5 = 180$. However, in some categories (e.g., artificial sweeteners, frozen fish, razor blades), there were fewer than five brands, which reduced the number of brands to 163. Furthermore, of these brands, 13 did not meet the minimum advertising frequency, which leaves us with 150 brands.

**Table 1**

<table>
<thead>
<tr>
<th>Product Class</th>
<th>Number of Categories</th>
<th>Example Categories</th>
<th>Example Brands</th>
</tr>
</thead>
<tbody>
<tr>
<td>Food</td>
<td>14</td>
<td>Breakfast cereals</td>
<td>Kellogg’s</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Canned fruits</td>
<td>Del Monte</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Yogurt</td>
<td>Danone</td>
</tr>
<tr>
<td>Beverages</td>
<td>7</td>
<td>Fruit juices and drinks</td>
<td>Ocean Spray</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Mineral water</td>
<td>Evián</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Soft drinks</td>
<td>Coca-Cola</td>
</tr>
<tr>
<td>Household care</td>
<td>7</td>
<td>Household cleaners</td>
<td>Mr. Muscle</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Machine wash products</td>
<td>Ariel</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Liquid detergents</td>
<td>Fairy</td>
</tr>
<tr>
<td>Personal care</td>
<td>8</td>
<td>Bath additives</td>
<td>Palmolive</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Toothpaste</td>
<td>Crest</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Shampoo</td>
<td>Head &amp; Shoulders</td>
</tr>
<tr>
<td>Total number</td>
<td>36</td>
<td></td>
<td>180</td>
</tr>
</tbody>
</table>
product category. Sixth, we need to account for possible endogeneity of price and advertising decisions. Subsequently, we formulate an error correction model that addresses these six challenges.² First, however, we present the derivation of the business cycle.

Assessing the State of the Business Cycle

To assess the impact of business cycles on marketing-mix effectiveness, we must first capture the business cycle itself. To do so, we adopt the Christiano–Fitzgerald (CF) random-walk filter (Christiano and Fitzgerald 2003). The CF filter is a band-pass filter built on the same principles as the Baxter and King filter used in Deleeuw et al. (2004). However, unlike the Baxter and King filter, which loses several observations at the end of the series, the CF filter is designed to use the entire time series (and thus also uses information from the most recent economic crisis).³ We refer to Christiano and Fitzgerald (2003) and Nilsson and Gyomai (2011) for a more in-depth discussion on the relative merits of the CF filter.

We apply the CF filter to log-transformed quarterly GDP data and transform the resulting series to a monthly sequence (the temporal aggregation level of our sales and marketing-support series) through linear interpolation (for similar practices, see Pauwels et al. 2004 and Srinivasan et al. 2009).⁴ The black line in Figure 2 shows the CF-filtered cyclical component (“Gross Domestic Product Cyclical”; gdpC). This is the cyclical deviation from the long-term trend in the log-transformed GDP series. The up- and downswing at the end of the data period reflects the strong performance of the U.K. economy in 2006 and 2007, followed by the global financial crisis that hit the United Kingdom in 2008.⁵ We categorize periods with an increase in the cyclical component as expansions and periods with a decrease as contractions (Lamey et al. 2007, 2012; Steenkamp and Fang 2011; see also Kontolemis 1997). In Figure 2, white zones represent expansions and gray zones represent contractions.⁶

Following Lamey et al. (2007) and Steenkamp and Fang (2011), we define the magnitude of the expansion at any point in time (Expansion_t) as the difference between the actual level of the cyclical component at time t and the prior trough. The prior trough is the most recent point in the cyclical component in which the month-on-month growth turned from negative to positive. Similarly, the extent of the contraction (Contraction_t) is the difference between current level and the prior peak (i.e., the most recent point in the cyclical component in which the month-on-month growth turned from positive to negative). The peaks and troughs also determine the transition points between the gray and white bars in Figure 2. Formally, we define the terms with the following equation:

\[
\text{Expansion}_t = \text{Cycle}_t - \text{Cycle}_{t-1} \\
\text{Contraction}_t = \text{Cycle}_{t-1} - \text{Cycle}_{t-2}
\]

²There are also other approaches to accommodate time-varying parameters, such as dynamic linear models (e.g., Ataman et al. 2010). However, estimating these models for our setting with its numerous brands, categories, and time periods would be prohibitive in terms of the amount of computer time it would take.

³Given that we work with quarterly data, a band-pass filter is more appropriate than the Hodrick–Prescott filter (which is a high-pass filter) used in other marketing applications (e.g., Deleeuw et al. 2009). Business-cycle fluctuations happen with a periodicity between 1.5 and 8 years (Christiano and Fitzgerald 2003). Researchers can design band-pass filters to pass through all components of the time series with a periodicity between 6 and 32 quarters. The Hodrick–Prescott filter, in contrast, suppresses only the low-frequency fluctuations, implying that what is passed through consists of both the cyclical fluctuations (of main interest) and higher-frequency noise. Although this does not make a difference when working with annual data (as in Deleeuw et al. 2009), it does have an impact when working with more aggregate (e.g., quarterly) data (for a more in-depth discussion, see Deleeuw et al. 2004, p. 357).

⁴Because the series is log-transformed before filtering, the resulting cyclical component (multiplied by 100) represents the percentage deviation from the economy’s underlying growth path (for a detailed discussion, see Lamey et al. 2012; Stock and Watson 1999).

⁵For a correct assessment of the last peak/trough preceding the 1993–2010 analysis period, we applied the CF filter to a data series that included several years before the analysis period.

⁶An alternative approach to identify contraction periods would be to use the recession periods, defined as at least two successive quarters of negative growth, as declared by such governmental agencies as the National Bureau of Economic Research Business Cycle Dating Committee (in the United States). However, researchers have criticized the committee’s judgment-based procedure for a lack of statistical foundation and for its rigid focus on absolute declines (as opposed to growth slowdowns) in output and other measures (Stock and Watson 1999). For these reasons, business-cycle filters have become the norm to study cyclical phenomena in contemporary business-cycle research.

Figure 2

(1) \[
\text{Expansion}_t = \begin{cases} 
\text{gdpc}_t - \left(\text{prior trough in gdpc}_t\right) & \text{if } \Delta\text{gdpc}_t > 0 \\
0 & \text{if } \Delta\text{gdpc}_t \leq 0
\end{cases}
\]

Contraction: \[
\begin{cases} 
0 & \text{if } \Delta\text{gdpc}_t > 0 \\
\left(\text{prior peak in gdpc}_t\right) - \text{gdpc}_t & \text{if } \Delta\text{gdpc}_t \leq 0
\end{cases}
\]

With this operationalization, Expansion, takes positive values during economic upswings and 0 values during downturns, whereas Contraction, takes positive values during downturns and 0 values during periods of growth. Stronger growth will result in higher values for Expansion, whereas a stronger decline will result in higher values for Contraction. By working with the continuous Expansion, and Contraction, variables, we not only account for the occurrence of the different business-cycle phases (as would be captured when working with dummy variables; see, e.g., Kontolemis 1997; Lamey et al. 2012) but also capture the intensity of the up- or downswing. Beauudy and Koop (1993) and Steenkamp and Fang (2011), among others, stress the importance of accounting for the current depth of the contraction/expansion. Many governmental agencies also define the depth of a recession relative to the previous trough (Siems 2012).

**Model for Assessing Price and Advertising Effectiveness over the Business Cycle**

Before specifying the model, we tested the (log-transformed) sales series for stationarity. Of the 150 individual sales series, the unit-root null hypothesis was rejected \((p < .10)\) in all but 7 (4.6%) instances on the basis of the Phillips–Perron (1988) test, using an intercept and trend as exogenous variables. Recent literature has pointed out that panel-based unit-root tests have higher power than tests based on the individual series. Therefore, we also applied Levin, Lin, and Chu’s (2002) test, allowing for both a fixed-effects correction and brand-specific trend, lag-lengths, and autoregressive parameters. Again, the unit-root null hypothesis was rejected \((p < .05)\) in favor of the no-unit-root alternative hypothesis. Therefore, the combined evidence of these tests supports the (trend) stationarity of brand sales, in line with other studies in the CPG industry (Dekimpe and Hanssens 1995a).

To capture the short- and long-term effects of advertising and price on brand sales, we adopt the parsimonious error correction specification (for recent marketing applications, see Fok et al. 2006; Pauwels, Srinivasan, and Franses 2007; Van Heerde, Helsen, and Dekimpe 2007; Van Heerde, Srinivasan, and Dekimpe 2010). For exposition purposes, we begin with a model with main effects only and subsequently add interactions with the business cycle:

(2) \[
\Delta\text{lnSales}^{cb}_t = \beta_0^{cb} \Delta\text{lnPrice}^{cb}_t + \beta_1^{cb} \Delta\text{lnAdvertising}^{cb}_t + \beta_2^{cb} \Delta\text{lnCompPrice}^{cb}_t + \beta_3^{cb} \Delta\text{lnCompAdvertising}^{cb}_t \]

\[
\Delta\text{lnSales}^{cb}_t = \beta_0^{cb} \Delta\text{lnPrice}^{cb}_t + \beta_1^{cb} \Delta\text{lnPrice}^{cb}_t \Delta\text{lnAdvertising}^{cb}_t + \beta_2^{cb} \Delta\text{lnPrice}^{cb}_t \Delta\text{lnCompPrice}^{cb}_t + \beta_3^{cb} \Delta\text{lnPrice}^{cb}_t \Delta\text{lnCompAdvertising}^{cb}_t + \text{lnSales}_{t-1}^{cb} - \beta_3^{cb} \Delta\text{lnPrice}_{t-1}^{cb} - \beta_2^{cb} \Delta\text{lnCompPrice}_{t-1}^{cb} - \beta_1^{cb} \Delta\text{lnAdvertising}_{t-1}^{cb} - \beta_0^{cb} \times \text{trend} + \epsilon_t^{cb}.
\]

where \(\Delta\) is the first difference operator: \(\Delta X_t = X_t - X_{t-1}\); \(\text{Sales}^{cb}_t\) is the volume sales of brand \(b\) \((b = 1 \ldots B_c)\) in category \(c\) \((c = 1 \ldots C)\) during month \(t\) \((t = 1 \ldots T)\), and \(ln\) indicates the natural logarithm. \(\text{Price}^{cb}_t\) is the deflated price of the brand at time \(t\), whereas \(\text{Advertising}^{cb}_t\) is the deflated advertising expenditure. We control for cross effects of competing brands’ actions by including the market-share-weighted average competitor price \(\text{CompPrice}_{t}^{cb}\) and total competitor advertising \(\text{CompAdvertising}^{cb}_t\).

In Model 2, \(\beta_1^{cb}\) is the short-term (same period) own price elasticity, \(\beta_2^{cb}\) is the short-term own advertising elasticity, \(\beta_3^{cb}\) is the short-term cross price elasticity, and \(\beta_4^{cb}\) is the short-term cross advertising elasticity; their long-term counterparts (same period + future periods) are \(\beta_5^{cb}, \beta_6^{cb}, \beta_7^{cb}\), and \(\beta_8^{cb}\), respectively. This long-term impact reflects the cumulative impact of a one-period shock to advertising or price, respectively. We refer to Fok et al. (2006) for an in-depth discussion of the error correction model. The term \(\beta_5^{cb}\) captures the long-term trend in brand sales as a proxy for other variables that gradually changed over the observed 18-year time span (Dekimpe and Hanssens 1995b; for a discussion of the role of this trend variable in error correction specifications, see Franses 2001). The trend variable runs from −1 in the first observation to +1 in the last. This enables us to interpret the main effect of the elasticity to hold for the middle observation (in which the trend variable = 0), and it makes the interactions and main effects of a similar magnitude. For the overall size and change in elasticity, it does not matter whether we define trend like this or, more traditionally, as running from 1 to T. For a discussion of the interpretational advantages of different coding schemes for trends, see Stoolmiller (1994). The parameter \(\varphi^{cb}\) reflects the speed of adjustment toward the underlying (long-term) equilibrium (Dekimpe and Hanssens 1999; Powers et al. 1991).

Equation 3 expands Equation 2 by adding the interactions between the own price and advertising variables and the Expansion and Contraction variables as well as the main effects of the latter variables:

(3) \[
\Delta\text{lnSales}^{cb}_t = \beta_0^{cb} + \beta_1^{cb} \Delta\text{lnPrice}^{cb}_t + \beta_2^{cb} \Delta\text{lnAdvertising}^{cb}_t + \beta_3^{cb} \Delta\text{lnPrice}^{cb}_t \times \text{Expansion}_t + \beta_4^{cb} \Delta\text{lnPrice}^{cb}_t \times \text{Contraction}_t + \beta_5^{cb} \Delta\text{lnPrice}^{cb}_t \times \text{trend}_t + \beta_6^{cb} \Delta\text{lnCompPrice}^{cb}_t + \beta_7^{cb} \Delta\text{lnCompPrice}^{cb}_t \times \text{Expansion}_t + \beta_8^{cb} \Delta\text{lnCompPrice}^{cb}_t \times \text{Contraction}_t + \beta_9^{cb} \Delta\text{lnCompPrice}^{cb}_t \times \text{trend}_t + \epsilon_t^{cb}.
\]

Model 3 is the base model to test the evolution of price and advertising elasticities over the business cycle. It puts the
focus squarely on the moderating role of the business cycle on own price and own advertising elasticities.

**Model Extensions**

We also consider three extensions to the base model. First, we allow for the possibility that own price elasticities or own advertising elasticities have changed over time. We do this by adding the four interactions of the trend with short- and long-term own elasticities to the base model: \( \Delta \text{lnPrice}_{cb}^t \times \text{trend} \), \( \Delta \text{lnAdvertising}_{cb}^t \times \text{trend} \), \( \text{lnPrice}_{c-1}^t \times \text{trend} \), and \( \text{lnAdvertising}_{c-1}^t \times \text{trend} \). Second, we investigate whether advertising moderates the effect of price by adding both the interaction between short-term own price and short-term own advertising (\( \Delta \text{lnPrice}_{cb}^t \times \Delta \text{lnAdvertising}_{cb}^t \)) and the interaction between long-term own price and long-term own advertising (\( \text{lnPrice}_{c-1}^t \times \text{lnAdvertising}_{c-1}^t \)) to Equation 3. Third, we consider whether the impact of short- and long-term effects of competitive advertising and competitive price varies across the business cycle, for a total of eight interaction terms. For example, to investigate the cyclical sensitivity of short-term cross price elasticities, we add \( \Delta \text{lnCompPrice}_{cb}^t \times \text{Expansion}_{cb} \) and \( \Delta \text{lnCompPrice}_{cb}^t \times \text{Contraction}_{cb} \). Adding all 14 interactions simultaneously to the base model leads to an overburdened model and unstable parameter estimates due to an excessive number of interaction terms (Cohen et al. 2003). Therefore, we examined each model extension separately by adding the terms to Equation 3 to evaluate their significance.

**Endogeneity**

Price and advertising decisions may depend on unobserved demand shocks. In addition, previous research has shown that not accounting for endogeneity may bias the elasticity estimates (Bijmolt, Van Heerde, and Pieters 2005; Sethuraman, Tellis, and Briesch 2011). We account for endogeneity by adopting a three-stage least squares approach (3SLS). The endogenous variables are \( \Delta \text{lnPrice}_{cb}^t \) and \( \Delta \text{lnAdvertising}_{cb}^t \) as well as any of the interactions of which they are a part. The lagged variables are predetermined.

As instrumental variables (IVs), we use lagged price and advertising variables from other product classes (not from the brand itself). We distinguish dairy food, nondairy food, beverages, household care, and personal care. So, for a beverage brand, we use as IVs average lagged log-price and average lagged log-advertising, separately for dairy food, nondairy food, household care, and personal care. We chose these IVs because the same underlying cost structures may determine price and advertising changes for other product classes. These cost structures are likely to be independent of the demand shocks observed for the focal product class. Lamey et al. (2012) and Ma et al. (2011), among others, recently adopted similar IVs in their studies. Because we also instrument for the interactions involving the first difference of price and advertising (e.g., \( \Delta \text{lnPrice}_{cb}^t \times \text{Expansion}_{cb} \)), the IVs also include the interactions between the original IVs and Expansion and Contraction (Wooldridge 2002, pp. 121–22). Because we have more IVs than endogenous regressors, our model is overidentified.

We formally test for both the strength (using the Angrist–Pischke [Angrist and Pischke 2009] multivariate F statistic) and the validity (using the Sargan test) of our instruments. These tests confirmed that our instruments are indeed correlated with the endogenous variables of interest (\( p \)-value F-tests < .05) but uncorrelated with the error term in the demand of the focal brand \( \varepsilon_{cb}^t \) (\( p \)-value Sargan test > .10).

**Model Estimation Procedure**

We use 3SLS to estimate the models, which takes into account error correlations between brands within each category. For the base model (Equation 3), we multiply the parameter \( \beta_{c}^{cb} \) by the term in the square brackets. Estimating the model, we initially obtain estimates for the different products of parameters (e.g., \( -\beta_{c}^{cb} \times \beta_{11}^{cb} \)), from which we can derive the estimate for the parameter of interest (e.g., \( \beta_{11}^{cb} \)). We subsequently derive the associated standard error with the Delta method (Greene 2000, p. 330–31).

We thus obtain brand-specific 3SLS estimates for all model parameters in Equation 3. To summarize the effect sizes and significance across all brands, we use Rosenthal’s method of added Zs (Rosenthal 1991). The effect size of parameter \( \beta \) is the weighted mean response parameter across brands. The weight is the inverse of the estimate’s standard error, normalized to one. Thus, we can interpret \( \beta \) as a reliability-weighted mean, where estimates with higher reliability (lower standard error) obtain a higher weight, similar in spirit to a hierarchical mean in a Bayesian model (e.g., Chib and Greenberg 1995).

We next add the interactions, one block at a time, and retain only the interactions that were significant at 10% across all brands, using the method of added Zs. Specifically, when we test the four interactions between trend and price and advertising (first difference and lag), the only term that is significant at 10% is \( \Delta \text{lnPrice}_{cb}^t \times \text{trend} \). When we test the short-term and long-term interactions between price and advertising, only the long-term interaction is significant: \( \Delta \text{lnPrice}_{c-1}^t \times \text{lnAdvertising}_{c-1}^t \). Finally, of the eight interactions between the various cross effects and the business cycle, only one is significant: \( \Delta \text{lnCompPrice}_{cb}^t \times \text{Contraction}_{cb} \). To avoid overspecifying the model, we retain only these three significant interactions (for a similar approach, see Bijmolt, Van Heerde, and Pieters 2005; Steenkamp and De Jong 2010). We call this the “extended model.”

**RESULTS**

**Overall Descriptive Findings**

Table 2 reports some descriptive statistics on the direction and extent to which the business cycle affects brand sales. This sets the stage for our main findings pertaining to the effect of the business cycle on our model parameters. Table 2 compares the average monthly sales during expansions (white areas in Figure 2) versus contractions (gray areas in Figure 2). Approximately the same number of CPG brands experienced a significant increase during a contraction as those that experienced a significant decrease. This may seem counterintuitive; however, it is consistent with Du and Kamakura’s (2008) finding that as discretionary income decreases, expenditures for essential categories (e.g., food at home) increase relative to nonessential cate-

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Footnote:

We refer to Web Appendix A (www.marketingpower.com/jmnr_webappendi-)

for more details on this procedure and to Deleersnyder et al. (2002,

2009) or Lamey et al. (2007, 2012) for recent marketing applications.
Table 2
CYCLICAL SENSITIVITY OF BRAND SALES

<table>
<thead>
<tr>
<th>Effect</th>
<th>Number of Brands</th>
<th>Typical Categories</th>
<th>Average Change</th>
</tr>
</thead>
<tbody>
<tr>
<td>Increased sales during contractions ($p &lt; .10$)</td>
<td>29</td>
<td>Conditioners</td>
<td>4%</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Cooking sauces</td>
<td>2%</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Pasta</td>
<td>8%</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Stout beer</td>
<td>6%</td>
</tr>
<tr>
<td>Decreased sales during contractions ($p &lt; .10$)</td>
<td>29</td>
<td>Household cleaner</td>
<td>−5%</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Liquid detergents</td>
<td>−2%</td>
</tr>
<tr>
<td>No significant effect</td>
<td>92</td>
<td>Shower products</td>
<td>−5%</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Cereal bars</td>
<td>−2%</td>
</tr>
<tr>
<td></td>
<td></td>
<td>Instant coffee</td>
<td>1%</td>
</tr>
<tr>
<td>Total number</td>
<td>150</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Table 2 also consistent with several recent observations in the business press that during tough economic times, consumers tend to switch from out-of-home consumption (e.g., restaurant, hairdresser) to more in-home consumption (Cendrowski 2012; Helm 2009; The Wall Street Journal Europe 2009). This favors grocery sales in those categories in which such a trading-in strategy is feasible, such as cooking sauces, pasta, stout beer, and hair conditioners. Conversely, restored consumer confidence during economic expansions tends to revive out-of-home consumption, putting pressure on grocery sales (Drake 2009).

Main Effects of Price and Advertising

Table 3 reports the parameter estimates for both the base model (Equation 3) and the extended model. Our discussion focuses on the results of the extended model. For parameters in which theory offers unidirectional predictions, we use one-sided $p$-values. For all other parameters, we use two-sided $p$-values. Table 3 includes a column with expectations for all model parameters.

We find a significant short-term ($\beta = -1.4266, p < .01$) and long-term price elasticity ($\beta = -.8379, p < .01$), which are consistent with the meta-analysis of Bijmolt, Van Heerde, and Pieters (2005) for CPG brands. We also find a significant short-term ($\beta = .0021, p < .10$) and long-term advertising elasticity ($\beta = .0127, p < .01$), both of which are small, but this is not uncommon in CPG categories (Allenby and Hanssens 2004; Sethuraman, Tellis, and Briesch 2011; Srinivasan, Vanhuele, and Pauwels 2010).

Is there evidence that price or advertising effectiveness has systematically declined or increased over time? For this, we consider the interactions between short- and long-term price and advertising elasticities with the trend. We find no evidence that advertising effectiveness declines over time. Although the parameter of both interaction terms is negative, indicating a tendency of declining advertising effectiveness over time, neither effect is significant ($p > .10$). However, price sensitivity has increased significantly over the past two decades. The relentless focus on price in the CPG industry has made consumers increasingly responsive to price reductions (Mela, Gupta, and Lehmann 1997; Mela, Jedidi, and Bowman 1998). The short-term price elasticity has become increasingly strong (i.e., more negative) over time, as implied by the significant interaction effect with the trend term ($\beta = -.5716, p < .01$). Because the trend variable runs from −1 to +1, this means that across the 18 years of data (1993–2010), the magnitude of the short-term price elasticity grew by 1.1432 percentage points. We find corroborating evidence by comparing our results with the meta-analysis of Bijmolt, Van Heerde, and Pieters (2005).

We observe that the annual change rate in the short-term price elasticity equals $2 \times (-.5716/18) = -0.0635$, which is close to the annual change rate of $-0.04$ reported by Bijmolt, Van Heerde, and Pieters (2005, p. 146).

Interaction Between Own Advertising and Own Price Effectiveness

We also examine whether price and advertising interact. From the notion that brand advertising should lead to increased product differentiation and reduced price sensitivity, we expect a positive interaction: more advertising should decrease the magnitude of the price elasticity (i.e., make it less negative). When we add the two interactions terms (for short-term and long-term elasticities, respectively) to the base model, we find that both have the expected positive sign, and the term for the long-term interaction is significant at 10%. However, in the extended model, which includes other interaction terms, this interaction effect is no longer significant. Therefore, we do not find compelling evidence that advertising effectiveness reduces price sensitivity across all brands and product categories.

Main Effect of the Business Cycle on Brand Sales

A correct interpretation of the effect of the business cycle parameters requires knowledge of the values of the two economy variables. Expansion, (Contraction,) shows an average value of 1.39 (1.45), with a minimum of 0 for both and a maximum of 5.31 (6.87), respectively. By multiplying the corresponding parameters with these values, we obtain an estimate of the size of the effect of the economy on the respective parameters.

We find that an expansion has a significant, negative main effect on brand sales ($\beta = -0.0103, p < .01$). This is consistent with the finding that during booms, the base sales of many CPG products will come under pressure because higher economic growth is associated with increased out-of-home consumption (Du and Kamakura 2008). To validate this finding, we calculated the correlation between the parameter estimate for the main effect of expansion on the intercept in the model (i.e., model-based evidence) and the percentage difference in average log sales between contractions and expansions (i.e., model-free evidence). The out-of-home argument suggests a negative correlation: the stronger the negative effect of expansion on the intercept, the larger (i.e., more positive) is the difference between average log sales in a contraction versus an expansion. In line with this argument, we find a correlation of $-0.25 (p < .01)$. The main effect of contractions is not significant, reflecting that the beneficial effects of contractions on some brands are balanced by the detrimental effects on other
Table 3
PARAMETER ESTIMATES

<table>
<thead>
<tr>
<th>Variable</th>
<th>Interpretation</th>
<th>Expectation</th>
<th>Base Model</th>
<th>Extended Model</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td></td>
<td>Weighted Mean $\hat{\beta}$</td>
<td>Meta-Analytic $Z$</td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td>Weighted Mean $\hat{\beta}$</td>
<td>Meta-Analytic $Z$</td>
</tr>
<tr>
<td>Intercept</td>
<td></td>
<td>$\neq 0$</td>
<td>.0747***</td>
<td>2.6715</td>
</tr>
<tr>
<td>Expansion</td>
<td>Main effect expansion on sales</td>
<td>$\neq 0$</td>
<td>-.0103***</td>
<td>-4.8699</td>
</tr>
<tr>
<td>Contraction</td>
<td>Main effect contraction on sales</td>
<td>$\neq 0$</td>
<td>.0004</td>
<td>.2346</td>
</tr>
<tr>
<td>$\Delta \ln Price^{cb}_t$</td>
<td>Short-term price elasticity</td>
<td>$&lt; 0$</td>
<td>-1.4355***</td>
<td>-20.9974</td>
</tr>
<tr>
<td>$\Delta \ln Price^{cb}_t \times Expansion_t$</td>
<td>Impact of expansion on short-term price elasticity</td>
<td>$&gt; 0$</td>
<td>-0.0284</td>
<td>-7.907</td>
</tr>
<tr>
<td>$\Delta \ln Advertising^{cb}_t$</td>
<td>Short-term advertising elasticity</td>
<td>$&gt; 0$</td>
<td>.0017*</td>
<td>1.4517</td>
</tr>
<tr>
<td>$\Delta \ln Advertising^{cb}_t \times Expansion_t$</td>
<td>Impact of expansion on short-term advertising elasticity</td>
<td>$\neq 0$</td>
<td>.0006</td>
<td>.8460</td>
</tr>
<tr>
<td>$\Delta \ln Advertising^{cb}_t \times Contraction_t$</td>
<td>Impact of contraction on short-term advertising elasticity</td>
<td>$\neq 0$</td>
<td>-0.0003</td>
<td>-7.113</td>
</tr>
<tr>
<td>$\Delta \ln Comp Advertising^{cb}_t$</td>
<td>Short-term cross advertising elasticity</td>
<td>$\neq 0$</td>
<td>.0011**</td>
<td>2.2799</td>
</tr>
<tr>
<td>$\Delta \ln Comp Price^{cb}_t$</td>
<td>Short-term cross price elasticity</td>
<td>$&lt; 0$</td>
<td>-.8117***</td>
<td>-18.1477</td>
</tr>
<tr>
<td>$\ln Price^{cb}_{t-1}$</td>
<td>Long-term price elasticity</td>
<td>$&gt; 0$</td>
<td>.0072*</td>
<td>1.3204</td>
</tr>
<tr>
<td>$\ln Price^{cb}_{t-1} \times Expansion_t$</td>
<td>Impact of expansion on long-term price elasticity</td>
<td>$&gt; 0$</td>
<td>-0.0048</td>
<td>-9.751</td>
</tr>
<tr>
<td>$\ln Price^{cb}_{t-1} \times Contraction_t$</td>
<td>Impact of contraction on long-term price elasticity</td>
<td>$&lt; 0$</td>
<td>.0012***</td>
<td>6.8236</td>
</tr>
<tr>
<td>$\ln Advertising^{cb}_{t-1}$</td>
<td>Long-term advertising elasticity</td>
<td>$&gt; 0$</td>
<td>.0018**</td>
<td>2.1602</td>
</tr>
<tr>
<td>$\ln Advertising^{cb}_{t-1} \times Expansion_t$</td>
<td>Impact of expansion on long-term advertising elasticity</td>
<td>$\neq 0$</td>
<td>-.0003</td>
<td>-4.485</td>
</tr>
<tr>
<td>$\ln Advertising^{cb}_{t-1} \times Contraction_t$</td>
<td>Impact of contraction on long-term advertising elasticity</td>
<td>$\neq 0$</td>
<td>.0085***</td>
<td>7.0154</td>
</tr>
<tr>
<td>$\ln Comp Price^{cb}_{t-1}$</td>
<td>Long-term cross advertising elasticity</td>
<td>$&gt; 0$</td>
<td>.4569***</td>
<td>9.5098</td>
</tr>
<tr>
<td>$\ln Sales^{cb}_{t-1}$</td>
<td>Long-term cross price elasticity</td>
<td>$&gt; 0$</td>
<td>.3905***</td>
<td>7.2529</td>
</tr>
<tr>
<td>Trend</td>
<td>Main effect trend</td>
<td>$\neq 0$</td>
<td>.1195***</td>
<td>11.2935</td>
</tr>
<tr>
<td>$\Delta \ln Price^{cb}_t \times trend$</td>
<td>Impact of trend on short-term price elasticity</td>
<td>$&lt; 0$</td>
<td>-.5716***</td>
<td>-4.5321</td>
</tr>
<tr>
<td>$\ln Price^{cb}<em>{t-1} \times ln Advertising^{cb}</em>{t-1}$</td>
<td>Long-term interaction advertising and price</td>
<td>$&gt; 0$</td>
<td>.0053</td>
<td>1.0897</td>
</tr>
</tbody>
</table>

*p < .10.
**p < .05.
***p < .01.

Notes: $p$-values are one-sided for directional hypotheses and two-sided otherwise. Weight for $\hat{\beta}$ is inverse of its standard error, normalized to 1.
brands. Correspondingly, the correlation between this main effect and the difference in average log-sales between contraction and expansion is an insignificant .027 (p = .743).

Moderating Role of the Business Cycle on Price and Advertising Elasticities

Although we find no evidence that the business cycle affects short-term elasticities, we find clear evidence that it affects the long-term impact of both price and advertising. It takes time for a firm to feel the full effects of economic ups and downs. Thus, myopic managers may well underestimate the full impact of economic contractions and expansions on the effectiveness of their marketing instruments. This is especially pertinent in economically difficult times, when managers are already under more stock market pressure to reach their (quarterly) targets (Deleersnyder et al. 2009; Steenkamp and Fang 2011) and when there is already less room for mistakes or suboptimal decisions.

More specifically, the interaction between long-term price elasticity and expansion is positive (β = .0097, p < .10). Conversely, long-term price elasticity becomes more negative during contractions (β = -.0084, p < .10). This implies that when the economy is doing well, consumers become less price sensitive, whereas price sensitivity increases in bad times (and more so the more severe the contraction).

Advertising elasticities increase significantly during expansions (β = .0017, p < .10), which means that demand becomes more responsive to advertising when the economy is in good shape. Our finding that the net impact of an expansion on advertising effectiveness is positive implies that in an expansion, the positive forces (e.g., consumers who are receptive to image advertising and more able to respond [Ang, Leong, and Kotler 2000; Sethuraman, Tellis, and Briesch 2011]) outweigh the negative forces (e.g., more clutter during expansions [Danaher, Bonfrer, and Dhar 2008]). Finally, advertising effectiveness is lower in bad economic times, though not significantly.

To get a sense of the magnitude of the cyclical swings in advertising and price elasticity, consider what is arguably the most important business cycle event in the post–World War II world, namely, the global financial crisis, which hit the United Kingdom in 2008 (see Figure 2). Figure 3 displays the long-term elasticities evaluated at the top of the peak in the economic cycle in November 2007 (indicated as “Expansion”) versus at the bottom of the trough (May 2009) of the global financial crisis (indicated as “Contraction”). The differences are substantial. The average long-term price sensitivity increases by 14% (from -.7863 to -.8956), whereas the advertising effectiveness drops approximately 60% (from .0218 to .0088).

Cross Brand Effects

What about the sensitivity of own brand sales to competitive price and advertising activity? Although our model controls for competitive prices and advertising in deriving own brand elasticities, cross brand elasticities are of managerial interest in their own right. Own brand sales are higher if the prices of competing brands increase. We find this effect both for the short run (β = .4803, p < .01) and the long run (β = .4128, p < .01). Both estimates are close to the meta-analytic average cross price elasticity of .52 in Sethuraman, Srinivasan, and Kim (1999). When we compare the magnitude of own versus cross price elasticities for the short run (−1.4266 vs. .4803) and the long run (−.8379 vs. .4128), it is clear that competitive pricing has a substantial effect on own brand sales. This is what we may expect in mature markets.8

Both the short-term (β = .0012, p < .05) and the long-term (β = .0086, p < .01) cross brand advertising elasticities are positive and significant. Thus, competitive advertising has a positive effect on own brand sales. Schultz and Wittink (1976) distinguish three advertising effects: primary sales effect (in which a brand’s advertising increases own sales without affecting competitive sales), primary demand effect (in which advertising increases own brand sales and those of its competitors), and competitive advertising (in which advertising increases own brand sales and decreases those of its competitors) (see also Hanssens, Parsons, and Schultz 2001, pp. 322–25). They use these effects to analytically derive six “cases of advertising effect” (Schultz and

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8In contrast, Steenkamp et al. (2005) documented relatively few significant cross brand effects. However, our analysis differs from their study on several dimensions, such as (1) the time span covered (4 years vs. 18 years), (2) the temporal aggregation level (weekly vs. monthly), (3) model specification (impulse-response functions derived from a vector auto-regression model vs. specific parameters in an error correction specification), and (4) level of entity aggregation (cross effects from individual brands vs. the cross effect of the combined (share-weighted) competition. Each of these factors can contribute to an observed difference in the obtained cross effects. An in-depth investigation of the relative contribution of each of these factors is beyond the scope of the current study.
Wittink, pp. 72–73). Given that we find that both own brand advertising and cross brand advertising increase sales and that own advertising elasticities are higher than cross elasticities, it appears that across all brands and categories, Schultz and Wittink’s (1976) “Case IV” describes the CPG market (at least in the United Kingdom). Case IV occurs if brand advertising positively affects sales of all brands, but has a stronger effect on the own brand.

There is little evidence that the magnitude of the cross elasticities systematically varies across the business cycle, though we acknowledge that the large number (eight) of possible interaction effects may have affected model stability (Cohen et al. 2003). However, we do find a significant effect of contractions on the long-term competitive price elasticity \((\beta = .0360, p < .05)\), implying that markets become more competitive when the going gets tough (Estelami, Lehmann, and Holden 2001).

### Cyclical Sensitivity of Price and Advertising Effectiveness Across Four Types of Brands

Thus far, we have focused on findings aggregated across a large number of brands. These findings provide input for empirical generalizations, which are much valued by marketing academics (Hanssens 2009). However, brand managers (and some marketing academics) may be more interested in findings for specific types of brands. Our model allows for heterogeneity in cyclical sensitivity of price and advertising effectiveness across brands.

To provide insights into the differences between brands, we categorize brands along two managerially relevant dimensions: their price level and their advertising support. From a median split, we obtain four types of brands: (1) high-advertising, high-price “premium mass brands” (e.g., Coca-Cola); (2) high-advertising, low-price “value mass brands” (e.g., Cif household cleaner); (3) low-advertising, high-price “premium niche brands” (e.g., Yorkshire Tea); and (4) low-advertising, low-price “value niche brands” (e.g., Bic razor blades). To avoid confounding brand and category characteristics, we carried out the median split within each category separately, with “low” < median value and “high” ≥ median value. We briefly discuss key differences between the four types of brands. Because the sample size in each cell is modest and differs between cells, we focus on relative magnitude of effects across the four types of brands rather than on significance levels. Web Appendix B (www.marketingpower.com/jmr_webappendix) reports detailed results, offering insights into which effects are significant and which are insignificant.

We focus on long-term effects to keep the discussion concise and because we have observed that the business cycle primarily affects long-term effectiveness of price and advertising. We again illustrate the effects of the business cycle using the global financial crisis. Figures 4 and 5 show the mean long-term price elasticity and mean long-term advertising elasticity, respectively, for each type of brand just before the crisis hit (labeled “Expansion”) versus the trough of the business cycle (“Contraction”).

On average, “premium mass brands” have comparatively high price elasticity and substantial advertising elasticity. However, the business cycle works out differently for these

![Figure 4](image-url)
two instruments. The direct effect of advertising on sales drops by about two-thirds in a contraction, but notably, price sensitivity hardly increases. Furthermore, premium mass brands are the only class of brands for which advertising substantially reduces the long-term price sensitivity through brand differentiation ($\beta = .0122$). Thus, it appears that for this brand type, high advertising buffers the negative price effect that would otherwise hit these mass brands.

Like premium mass brands, “value mass brands” exhibit comparatively high price sensitivity, but unlike the former, advertising does not buffer price sensitivity in contractions. The interaction between price and advertising is negligible, and these brands exhibit the largest absolute increase in price sensitivity in contractions. Figure 4, Panel B, illustrates this: the long-term price elasticity goes from $-0.7735$ at the peak of the expansion to $-1.0035$ at the bottom of the contraction. Advertising is not particularly effective, and its relative magnitude declines substantially in a contraction. Sales of value mass brands are also much more vulnerable to competitive pricing than premium mass brands (long-term cross price elasticity is .4610 and .2141, respectively). To maintain sales in a contraction, value mass brands will have little alternative but to resort to price competition. This will be a tough battle, because value brand prices are already low.

Niche brands exhibit lower price sensitivity than mass brands. Indeed, they have an advantage in that their downward potential is dampened because some consumers prefer their characteristics no matter what. On the downside, their upward potential is also limited because a lot of people supply are not interested in the brand proposition. But this is where the similarity between premium niche and value niche brands ends. “Premium niche brands” find it difficult to justify the price premium in an economic decline, with a large percentage increase in price elasticity during contractions (31%, albeit from a lower basis). Moreover, the cross price elasticity increases substantially in contractions ($\beta = .1014$). Managers cannot employ advertising to counter price pressure because it essentially has no effect on sales during contractions (see Figure 5, Panel C). However, these brands are still in a better position than value mass brands. Their price sensitivity is lower and their higher price gives them more room to reduce prices in tough times.

Finally, though “value niche brands” do not advertise often, it is rather effective when they do. Advertising effectiveness declines substantially in contractions, but it can still offer an antidote to the negative effects of increased price sensitivity. Regardless of the instrument used, however, additional stress will affect an already-strained bottom line, as either the margin is further reduced (in case of stronger price reduction) or a higher increase in fixed costs (in case of advertising hikes) is called for.

**Cyclical Sensitivity of Price and Advertising Effectiveness Across Product Classes**

Whereas the cyclical sensitivity of price and advertising effectiveness is important for brand managers, senior managers (typically [vice] presidents) who are responsible for entire business units may be especially interested in the cyclical sensitivity of broad product classes. In this section,
we provide results for four main types of CPG products: (1) food, (2) beverages, (3) household care, and (4) personal care. We again focus on the long-term elasticities. In Figures 6 and 7, we show how the effectiveness of advertising and price changes from the expansion days near the end of 2007 to the trough around mid-2009.

**Beverages** show the strongest overall impact of the business cycle. Their price sensitivity increases 27% (from \(-0.9483\) to \(-1.2096\)), and their advertising becomes largely ineffective (−76%, from .0251 to .0059). The decline in advertising effectiveness is not buffered by a moderating effect of advertising on price sensitivity, which is the smallest of all product classes (\(\beta = .0004\)). Furthermore, the long-term cross price elasticity (\(\beta = .6668\)) is higher than in any other product class. The message is clear: in steep contractions, beverage brands need to stay competitive on price. If the goal is to protect brand sales, the person responsible for the beverages division may consider shifting money from advertising to pricing.

The business cycle does not affect **food** brands nearly as much as is does beverage brands. Food brands become somewhat more price sensitive (increasing from \(-0.8786\) to \(-1.0101\), a 14% change), but their advertising effectiveness is hardly affected. In this product class, it makes sense to remain price competitive during contractions (after all, the price elasticity is substantial) and to invest in advertising. In bad times, one witnesses a shift from out-of-home to in-home consumption. In the words of investment fund manager Jeff Auxier, “People will eat more meals at home....They do not eat dinner out, but they’ll still buy Ben & Jerry’s for dessert” (Cendrowski 2012, p. 76). Clearly, in such conditions, advertising helps to keep the brand top-of-mind and to communicate core values.

The business cycle has little effect on the price sensitivity of **household care** brands, but advertising effectiveness is nearly cut in half, although it remains relatively effective. Furthermore, the cross price elasticity is lowest of all product classes (\(\beta = .2299\)). In contractions, companies that are active in multiple product classes (e.g., Unilever, Procter & Gamble) may consider shifting marketing dollars from household care brands to brands in other product classes, particularly food brands, which have higher price and advertising elasticities in contractions, and possibly beverage brands, which need more price support in contractions.

Finally, the price sensitivity of **personal care** brands is relatively low, and the business cycle hardly affects it. In favorable times, the direct effect of advertising on sales is low (see Figure 7, Panel D), but personal care is the only product class in which there is an appreciable interaction between advertising and price (\(\beta = .0099\)). This indicates that by reducing pricing sensitivity, advertising has a secondary effect on sales. In contractions, advertising’s direct effect becomes negative. As we discussed previously, Ang, Leong, and Kotler (2000) argue that in contractions, consumers focus less on brands’ image aspects (which are relatively more important in personal care) and, moreover, that brands relying heavily on advertising may be viewed as being less sympathetic to the consumer’s tight economic situation. This can explain why we find that for personal care brands, advertising can indeed be in danger of becoming counterproductive (note that it still exerts an indirect positive effect through price sensitivity). We further find that the

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**Figure 6**

THE EFFECT OF THE GLOBAL FINANCIAL CRISIS: LONG-TERM PRICE ELASTICITY FOR FOUR PRODUCT CLASSES

<table>
<thead>
<tr>
<th>A: Food</th>
<th>B: Beverages</th>
</tr>
</thead>
<tbody>
<tr>
<td>Expansion</td>
<td>Contraction</td>
</tr>
<tr>
<td>(-0.88)</td>
<td>(-1.01)</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>C: Household Care</th>
<th>D: Personal Care</th>
</tr>
</thead>
<tbody>
<tr>
<td>Expansion</td>
<td>Contraction</td>
</tr>
<tr>
<td>(-0.97)</td>
<td>(-0.96)</td>
</tr>
</tbody>
</table>
cross price elasticity strongly increases in economic downturns (.1379). Thus, in contractions, personal care brands might consider cutting back on advertising and reserve money to respond more vigorously to rival pricing moves.

DISCUSSION

Summary

Although marketing effectiveness has been the subject of a wide stream of research, to the best of our knowledge, this is the first study that investigates how general economic conditions may influence the effectiveness of both brand advertising and brand pricing, using a single-source dedicated data set for many brands over a long period of 18 years encompassing multiple expansions and contractions, including the global financial crisis. Our single-source data enable us to use a consistent measurement and modeling scheme across all brands, which eliminates the confounding factors identified by earlier meta-analyses on price and advertising, such as differences in measures, temporal intervals, aggregation, model specification, estimation method, endogeneity, and heterogeneity (Bijmolt, Van Heerde, and Pieters 2005; Sethuraman, Tellis, and Briesch 2011).

We estimate an error correction model that provides both short- and long-term price and advertising elasticities, accommodates potentially asymmetric effects of the business cycle on these elasticities, and allows for heterogeneity between brands. We account for the possible endogeneity of price and advertising decisions by using a 3SLS approach. We investigate the magnitude of short- and long-term price and advertising elasticities, their interaction and evolution over time, and most pertinently, how price and advertising brand elasticities vary across the business cycle. Furthermore, we examine to what extent the cyclical sensitivity of advertising and price differs for four types of brands (premium mass brands, value mass brands, premium niche brands, and value niche brands) and four product classes (food, beverages, household care, and personal care).

Our findings show that although the short-term price and advertising elasticities do not change with the business cycle, the long-term elasticities do. The long-term price sensitivity decreases during expansions, whereas it increases during economic downturns. Contractions also make the long-term competitive price effects stronger. The long-term advertising elasticity, in turn, becomes stronger during economic expansions. Finally, patterns are not symmetric across contraction and expansion periods and differ systematically between premium mass brands, value mass brands, premium niche brands, and value niche brands as well as across major product classes. We also find that across the 18 years we observed, consumers’ short-term price elasticity has gradually grown in magnitude, whereas advertising effectiveness has not declined appreciably.

Managerial Implications

Firms are under increasing pressure to improve both the accountability and the effectiveness of their marketing investments. Our findings can guide managers in choosing strategies when deciding marketing investments to build brand sales across the business cycle. During contractions, managers are under especially close scrutiny as to how they...
allocate their budgets over price reductions (by lowering the margins) and advertising. This article offers answers to this quandary. Compared with expansions, during contractions, consumers are less responsive to advertising; at the same time, they react more strongly to price reductions. Application of the Dorfman–Steiner rule (Dorfman and Steiner 1954) leads us to recommend reallocation of budgets from advertising to price discounts in tough times.

However, this recommendation is not without several caveats. First, it is made in the context of maintaining brand sales. Firms can have other objectives, such as maintaining profits or securing shelf space. Second, there is considerable heterogeneity across product classes in the cyclical sensitivity of their long-term price and advertising elasticities. For example, advertising elasticity remains comparatively high in contractions for food brands, but not so for beverages. Third, there is also considerable heterogeneity depending on the brand’s positioning within the category. Premium mass brands, such as Gillette, Kellogg’s, or Coca-Cola, present an interesting case. Whereas advertising’s direct effect on sales declines dramatically in contractions, advertising has an important indirect effect on sales by reducing price sensitivity. Therefore, in these cases there is less reason to follow a procyclical advertising strategy and lower prices during downturns than for other brands. As such, we identify another way through which advertising may enhance brand performance: it not only increases the willingness to pay (Steenkamp, Van Heerde, and Geyskens 2010) and lowers the price sensitivity (Ataman, Van Heerde, and Mela 2010), it also insulates a brand’s marketing effectiveness from the impact of the business cycle (as we illustrate in this study).

We now encounter an important question: how can we reconcile our finding that the long run-advertising elasticity evolves in a procyclical way with observations that companies spending relatively more on advertising during economic downturns have better financial performance (Deleeuw–snyder et al. 2009; Steenkamp and Fang 2011)? We believe the answer lies in the nature of CPG products. We find that, across the board, CPG sales move countercyclically. When the economy is expanding, CPG sales tend to suffer because consumers engage in more out-of-home consumption. Thus, while the long-term advertising elasticity moves procyclically relative to the overall economy, it moves in a countercyclical way compared with CPG sales. Therefore, an expansion means relatively bad times for CPG sales, and it is then that advertising becomes more effective in the long run. This is consistent with the idea that spending during hard times (for the focal industry) is beneficial.

Our results show that economic downturns should not necessarily be detrimental for CPG companies, given the insignificant effect of contractions on base sales. The Boston Consulting Group provides additional evidence for these findings, noting that 58% of companies that were among the top three in their industry had rising profits in 2008, and only 30% saw their profits decline (The Economist 2009). Although danger signs are all around during economic downturns, CPG companies should grasp the opportunities they represent as well. In terms of their base sales, CPG companies may be able to capitalize on the positive substitution effect from out-of-home consumption during economic downturns. However, their higher vulnerability to competitive price attacks and the reduced effectiveness of some of their own instruments indicate that economic downturns are no easy ride for those managers. Depending on a company’s product class and its relative positioning within that class, opportunities exist to ride the economic tides profitably. Our results help identify which strategic adjustments are more or less opportune for different brands and product classes.

Directions for Further Research

Our study has several limitations that offer opportunities for further research. First, we excluded private labels, because their marketing support is typically governed by very different decision processes than national brands. Even though some retailers have recently begun to advertise their private labels, their ad intensity is minor compared with national brands—if present, it is almost never product specific (Corstjens and Steele 2008; Lamey et al. 2012). However, examining to what extent marketing-mix effectiveness varies over the business cycle for private labels is an interesting avenue for further research, especially given these labels’ remarkable and persistent market-share gains during contractions (Lamey et al. 2007).

Second, we based our analyses on relatively mature CPG categories. Such products are characterized by very small advertising elasticities. Conversely, we expect less mature categories to show stronger advertising sensitivity (Allenby and Hanssens 2004). In addition, this sensitivity may also vary more with the overall economic sentiment. Future studies could include such products. In addition, new products are known to have higher advertising effectiveness. Because we focused on major brands that had been in the market for a long time, the question remains whether economic downturns equally affect more recent brands.

Third, the products in our data set are mainly everyday consumables, purchases that cannot really be postponed until the economy recovers. This is not the case for durables: consumers can and do wait to purchase until economic conditions improve and uncertainty diminishes (Deleersnyder et al. 2004). This could result in even stronger business-cycle influences on marketing-mix effectiveness. Therefore, we call for a deeper investigation into this issue.

Finally, given the nature of our data, we estimated an aggregate response model. Although it would be difficult to obtain individual-level data from a representative panel across multiple business cycles, estimating brand-choice models in the tradition of Seetharaman, Ainslie, and Chintagunta (1999) is likely to result in additional insights. Indeed, this would enable studying how consumers (segments) react differently, in terms of changes in their advertising and price responsiveness, to changing economic conditions.

In summary, this study provides insights on the evolution of price and advertising effectiveness across the business cycle over a large set of CPG brands. We hope that this research inspire additional studies that will help brand managers to better ride the economic tides.

REFERENCES


